

Random or stochastic processes

Definition

- Random experiment specified by $\zeta \in S$, by the events defined on S , and by the probabilities on these events. To each $\zeta \in S$, a function of time is assigned according to some rule.

$$X(t, \zeta) \quad t \in I$$

$X(t, \zeta) \Big|_{\zeta}$ is a function of time t
realization, sample path, sample function

$X(t, \zeta) \Big|_{t=t_k}$ is a function of ζ , i.e. an r.v.

$\{X(t, \zeta), t \in I\}$ an indexed family of r.v.'s
or random process
or stochastic process



Ex 6.1 discrete-time SP

$$\zeta \in S = (0, 1]$$

$$\forall \zeta \in S \rightarrow \zeta = \sum_{n=1}^{\infty} b_n 2^{-n} \quad \text{binary expansion}$$

discrete-time SP: $X(n, \zeta) = b_n ; n = 1, 2, \dots$

Ex 6.2 continuous-time SP

$$\zeta \in S = [-1, 1]$$

$$X(t, \zeta) = \zeta \cos(2\pi t); -\infty < t < \infty$$

sinusoids with amplitude ζ

$$\zeta \in S = [-\pi, \pi]$$

$$Y(t, \zeta) = \cos(2\pi t + \zeta); -\infty < t < \infty$$

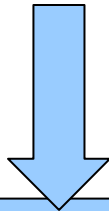
time-shifted versions of cos

Specifying a SP

sample the SP $X(t, \zeta) \Big|_{t=t_i} \rightarrow$ r.v.'s X_1, X_2, \dots, X_k

joint behavior of the SP at these k time instants is specified by the joint CDF of the vector r.v. $\mathbf{X} = (X_1, X_2, \dots, X_k)$

we can now compute probabilities as before



a SP is specified by the collection of k^{th} -order joint CDFs:

$$F_{X_1, \dots, X_k}(x_1, \dots, x_k) = P[X_1 \leq x_1, X_2 \leq x_2, \dots, X_k \leq x_k]$$

for any k and any choice of sampling instants t_1, \dots, t_k

$$p_{X_1, \dots, X_k}(x_1, \dots, x_k) = P[X_1 = x_1, X_2 = x_2, \dots, X_k = x_k] \quad \text{for discrete-time SP}$$

$$f_{X_1, \dots, X_k}(x_1, \dots, x_k) \quad \text{for continuous-time SP}$$

partial characterizations using moments

Mean, autocorrelation, autocovariance

$$m_X(t) = E[X(t)] = \int_{-\infty}^{\infty} x f_{X(t)}(x) dx$$

$$R_X(t_1, t_2) = E[X(t_1)X(t_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{X(t_1)X(t_2)}(x, y) dx dy$$

$$\begin{aligned} C_X(t_1, t_2) &= E\left[\{X(t_1) - m_X(t_1)\}\{X(t_2) - m_X(t_2)\}\right] \\ &= R_X(t_1, t_2) - m_X(t_1)m_X(t_2) \end{aligned}$$

$$VAR[X(t)] = E\left[\{X(t) - m_X(t)\}^2\right] = C_X(t, t)$$

$$\rho_X(t_1, t_2) = \frac{C_X(t_1, t_2)}{\sqrt{C_X(t_1, t_1)C_X(t_2, t_2)}} \implies |\rho_X(t_1, t_2)| \leq 1$$



Ex 6.6 random amplitude sinusoid

$$X(t) = A \cos 2\pi t$$

$$m_X(t) = E[A \cos 2\pi t] = E[A] \cos 2\pi t$$

mean varies with t

process is always =0 for t -values
for which $\cos 2\pi t = 0$

$$\begin{aligned} R_X(t_1, t_2) &= E[A \cos(2\pi t_1) A \cos(2\pi t_2)] \\ &= E[A^2] \cos(2\pi t_1) \cos(2\pi t_2) \end{aligned}$$

$$\begin{aligned} C_X(t_1, t_2) &= R_X(t_1, t_2) - m_X(t_1) m_X(t_2) \\ &= \{E[A^2] - E[A]E[A]\} \cos(2\pi t_1) \cos(2\pi t_2) \\ &= \text{VAR}(A) \cos(2\pi t_1) \cos(2\pi t_2) \end{aligned}$$



Ex 6.7 random phase sinusoid

$$X(t) = \cos(\omega t + \Theta); \Theta \sim U(-\pi, \pi)$$

$$m_X(t) = E[\cos(\omega t + \Theta)] = \frac{1}{2\pi} \int_{-\pi}^{\pi} \cos(\omega t + \theta) d\theta = 0$$

mean is constant

$$\begin{aligned} C_X(t_1, t_2) &= R_X(t_1, t_2) = E[\cos(\omega t_1 + \Theta) \cos(\omega t_2 + \Theta)] \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{1}{2} \{ \cos(\omega(t_1 - t_2)) + \cos(\omega(t_1 + t_2) + 2\theta) \} d\theta \\ &= \frac{1}{2} \cos(\omega(t_1 - t_2)) \end{aligned}$$

covariance depends on absolute time difference only



Gaussian SP

A random process $X(t)$ is a **Gaussian SP**

if for any k and any choice of sampling instants $t_1 < t_2 < \dots < t_k$

the r.v.'s $X_1 = X(t_1), X_2 = X(t_2), \dots, X_k = X(t_k)$

are jointly Gaussian r.v.'s

discrete & continuous

$$\begin{aligned} f_{\mathbf{X}}(\mathbf{x}) &\triangleq f_{X_1, X_2, \dots, X_k}(x_1, x_2, \dots, x_k) \\ &= \frac{\exp\left\{-\frac{1}{2}(\mathbf{x} - \mathbf{m})^T K^{-1}(\mathbf{x} - \mathbf{m})\right\}}{(2\pi)^{k/2} |K|^{\frac{1}{2}}} \end{aligned}$$



k jointly Gaussian r.v.'s

r.v.'s X_1, X_2, \dots, X_k are said to be jointly Gaussian if their joint PDF is given by

$$f_{\mathbf{X}}(\mathbf{x}) \triangleq f_{X_1, X_2, \dots, X_k}(x_1, x_2, \dots, x_k) \\ = \frac{\exp\left\{-\frac{1}{2}(\mathbf{x} - \mathbf{m})^T K^{-1}(\mathbf{x} - \mathbf{m})\right\}}{(2\pi)^{k/2} |K|^{\frac{1}{2}}}$$

$$\mathbf{m} = \begin{bmatrix} m_X(t_1) \\ m_X(t_2) \\ \vdots \\ m_X(t_k) \end{bmatrix} \quad K = \{C_X(t_i, t_j)\}_{i,j}$$



Multiple SPs

- Joint behavior of two or more SPs is specified by the collection of joint distributions for all possible choices of time samples of the SPs
- e.g. for $X(t)$ and $Y(t)$ specify all possible joint density functions of $X(t_1), \dots, X(t_k)$ and $Y(t_1'), \dots, Y(t_j')$ for all k, j and all choices of t_1, \dots, t_k and t_1', \dots, t_j'



Multiple SPs

- The SPs $X(t)$ and $Y(t)$ are said to be **independent** if the vector r.v.'s $(X(t_1), \dots, X(t_k))$ and $(Y(t_1'), \dots, Y(t_j'))$ are independent for all k, j and all choices of t_1, \dots, t_k and t_1', \dots, t_j'

cross-correlation $R_{XY}(t_1, t_2) = E[X(t_1)Y(t_2)]$

SPs $X(t)$ and $Y(t)$ are said to be **orthogonal** if $R_{XY}(t_1, t_2) = 0 \forall t_1, t_2$

cross-covariance

$$\begin{aligned} C_{XY}(t_1, t_2) &= E\left[\{X(t_1) - m_X(t_1)\}\{Y(t_2) - m_Y(t_2)\}\right] \\ &= R_{XY}(t_1, t_2) - m_X(t_1)m_Y(t_2) \end{aligned}$$

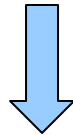
SPs $X(t)$ and $Y(t)$ are said to be **uncorrelated** if $C_{XY}(t_1, t_2) = 0 \forall t_1, t_2$



Ex 6.10 signal + noise

$$Y(t) = X(t) + N(t)$$

cross-correlation



$$\begin{aligned} R_{XY}(t_1, t_2) &= E[X(t_1)Y(t_2)] \\ &= E[X(t_1)\{X(t_2) + N(t_2)\}] \\ &= E[X(t_1)X(t_2)] + E[X(t_1)N(t_2)] \\ &= R_X(t_1, t_2) + E[X(t_1)]E[N(t_2)] \\ &= R_X(t_1, t_2) + m_X(t_1)m_N(t_2) \end{aligned}$$

independence



Stationary random processes

- Essentially...the nature of the randomness does not change with time
- Assuming the process started at $t=-\infty$...
A discrete-time or continuous-time SP $X(t)$ is stationary if the joint distribution of any set of samples does not depend on the placement of the time origin

$$F_{X(t_1), \dots, X(t_k)}(x_1, \dots, x_k) = F_{X(t_1+\tau), \dots, X(t_k+\tau)}(x_1, \dots, x_k); \forall \tau, \forall k, \forall t_1, \dots, t_k$$

if SP starts at definite time, eg. $t=0$ or $n=0$, joint CDF does not change under time shifts to the right



Jointly stationary SPs

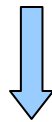
- Two processes $X(t)$ and $Y(t)$ are said to be jointly stationary if

$$\begin{aligned} F_{X(t_1), \dots, X(t_k), Y(t'_1), \dots, Y(t'_j)}(x_1, \dots, x_k, y_1, \dots, y_j) &= \\ &= F_{X(t_1+\tau), \dots, X(t_k+\tau), Y(t'_1+\tau), \dots, Y(t'_j+\tau)}(x_1, \dots, x_k, y_1, \dots, y_j) \\ &\quad \forall \tau, \forall k, j, \forall t_1, \dots, t_k, \forall t'_1, \dots, t'_j \end{aligned}$$



First-order CDF of stationary SP

$$\begin{aligned} F_{X(t)}(x) &= F_{X(t+\tau)}(x); \forall \tau, \forall t \\ &= F_X(x) \end{aligned}$$



$$m_X(t) = E[X(t)] = m; \forall t$$

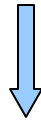
$$\text{VAR}(X(t)) = E[(X(t) - m)^2] = \sigma^2; \forall t$$

mean and variance are constant



Second-order CDF of stationary SP

$$\begin{aligned} F_{X(t_1), X(t_2)}(x_1, x_2) &= F_{X(t_1+\tau), X(t_2+\tau)}(x_1, x_2) \\ &= F_{X(0), X(t_2-t_1)}(x_1, x_2); \forall t_1, t_2 \end{aligned}$$



$$R_X(t_1, t_2) = R_X(t_2 - t_1); \forall t_1, t_2$$

$$C_X(t_1, t_2) = C_X(t_2 - t_1); \forall t_1, t_2$$

autocorrelation and covariance depend only on time difference



Wide-sense stationary (WSS) SP

- Often we cannot determine whether a process is stationary, but we can determine whether

- *the mean is constant* $m_X(t) = m; \forall t$

- *the autocovariance (equivalently the autocorrelation) is a function of time difference only* $C_X(t_1, t_2) = C_X(t_1 - t_2); \forall t_1, t_2$

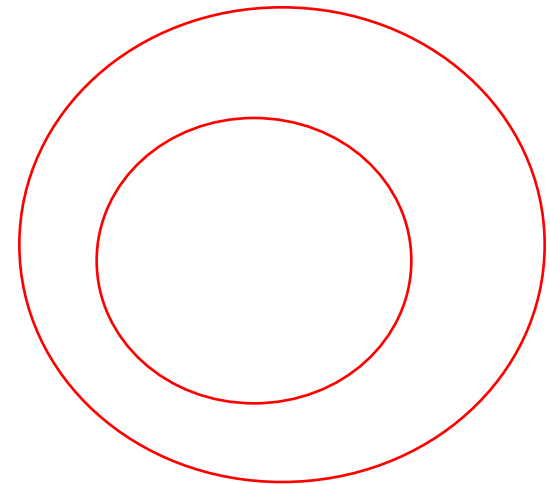
- SP is WSS if **both** conditions hold



Jointly WSS SPs

- $X(t)$ and $Y(t)$ are jointly WSS SPs if
 - $X(t)$ is WSS
 - $Y(t)$ is WSS
 - $C_{XY}(t_1, t_2) = C_{XY}(t_1 - t_2) = C_{XY}(\tau)$

- All stationary SPs are WSS



Autocorrelation function of WSS SP

- Plays a crucial role in the design of linear signal processing algorithms

i $R_X(0) = E[X^2(t)]$ average power of the process

ii $R_X(\tau) = E[X(t+\tau)X(t)] = E[X(t)X(t+\tau)] = R_X(-\tau)$
an even function of τ

iii $E^2[XY] \leq E[X^2]E[Y^2] \forall X, Y$

$$R_X^2(\tau) = E^2[X(t+\tau)X(t)] \leq E[X^2(t+\tau)]E[X^2(t)] = R_X^2(0)$$

$$|R_X(\tau)| \leq R_X(0)$$

autocorrelation is max at $\tau=0$

- If $R_x(d) = R_x(0) \rightarrow R_x(\tau)$ is periodic with period equal to d .

- Let $X(t) = m + N(t)$

$N(t)$ is a zero-mean process $\ni R_N(\tau) \xrightarrow{\tau \rightarrow \infty} 0$

$$R_X(\tau) = E\left[\{m + N(t + \tau)\}\{m + N(t)\}\right]$$

$$= m^2 + 2mE[N(t)] + R_N(\tau)$$

$$= m^2 + R_N(\tau) \xrightarrow{\tau \rightarrow \infty} m^2$$

$R_X(\tau)$ approaches m^2 (the square of the mean) as $\tau \rightarrow \infty$



Autocorrelation function of WSS SP

- Can have three types of component:
- A component that vanishes as $\tau \rightarrow \infty$
- A periodic component
- A component due to a nonzero mean



Ex 6.30

$$R_X(\tau) = e^{-2\alpha|\tau|}$$

$X(t)$ is zero-mean and $R_X(\tau) \xrightarrow{\tau \rightarrow \infty} 0$

$$R_Y(\tau) = \frac{a^2}{2} \cos(2\pi f_0 \tau) \quad \text{random phase sinusoid}$$

$Y(t)$ is zero-mean and $R_X(\tau)$ has period f_0^{-1}

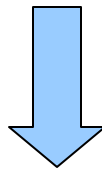
$$R_Z(\tau) = R_X(\tau) + R_Y(\tau) + m^2$$

$Z(t) = X(t) + Y(t) + m$; $X(t)$ and $Y(t)$ independent

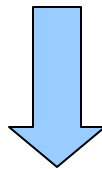
Note that it is not possible to determine which fraction of the mean is contributed by the individual components

WSS Gaussian SPs

- Mean is not a function of t
- Autocovariance is a function of time difference only

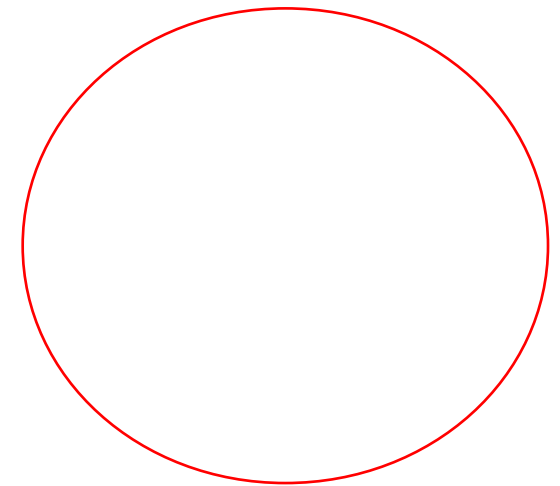


- Joint PDF depends on mean and autocovariance only



- A WSS Gaussian SP is stationary

all we need is m and $C_X(\tau)$





Ex 6.31

X_n is iid sequence of $N(0, \sigma^2)$ r.v.'s. Y_n is the average of two consecutive values

$$Y_n = \frac{X_n + X_{n-1}}{2} \longrightarrow m_Y = 0.5E[X_n + X_{n-1}] = 0$$

$$\begin{aligned} C_Y(i, j) &= E[Y_i Y_j] = 0.25E[(X_i + X_{i-1})(X_j + X_{j-1})] \\ &= 0.25\{E[X_i X_j] + E[X_i X_{j-1}] + E[X_{i-1} X_j] + E[X_{i-1} X_{j-1}]\} \\ &= \frac{1}{2}\sigma^2\delta_{i-j} + \frac{1}{4}\sigma^2\delta_{i-j+1} + \frac{1}{4}\sigma^2\delta_{i-j-1} \quad Y_n \text{ is WSS} \end{aligned}$$

Y_n is a linear trafo of Gaussian r.v.'s, so Y_n is Gaussian

joint PDF is Gaussian, specified by m_Y and $C_Y(i, j)$



Time averages & ergodicity

$$\hat{m}_X(t) = \frac{1}{N} \sum_{i=1}^N X(t, \zeta_i)$$

ensemble averaging
repeating the experiment
from many realizations

$$\langle X(t) \rangle_T = \frac{1}{2T} \int_{-T}^T X(t, \zeta) dt$$

time averaging
based on single realization



an ergodic theorem states conditions under which a time average converges when the observation interval becomes large

we're interested in ergodic theorems that state when time averages converge to the ensemble average or expected value

$$X(t) = A \cos(2\pi f_c t + \vartheta)$$

ϑ is U. RV from $0 \longrightarrow 2\pi$

Is $X(t)$ mean and autocorrelation ergodic?