

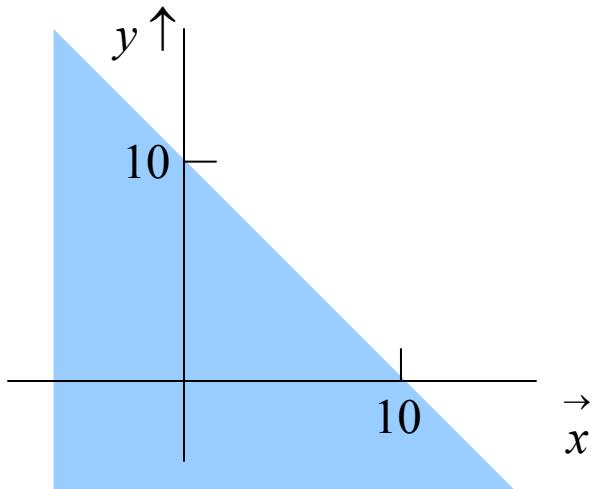
Multiple r.v.'s

- Several r.v.'s at a time
 - *Measuring different quantities simultaneously*
 - Engine oil pressure, RPM, generator voltage
 - *Repeated measurement of the same quantity*
 - Sampling a waveform, such as EEG, speech
- Joint behavior of two or more r.v.'s
 - *Independence of sets of r.v.'s*
 - *Correlation if not independent*

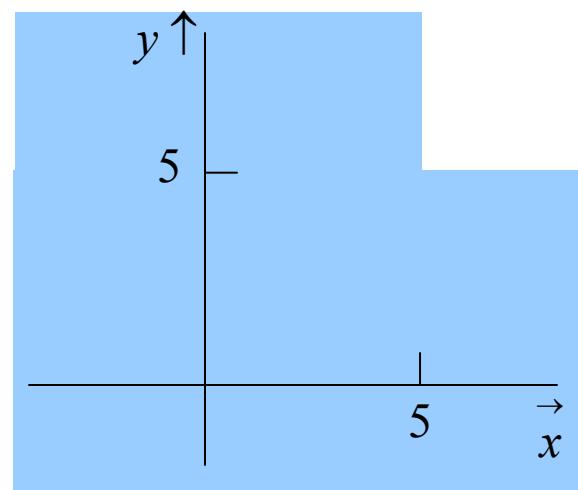
Events & probabilities

- Each event involving an n -dimensional r.v. $\mathbf{X}=(X_1, X_2, \dots, X_n)$ has a corresponding region in an n -dimensional real space
 - E.g. 2-D r.v. $\mathbf{X}=(X, Y)$

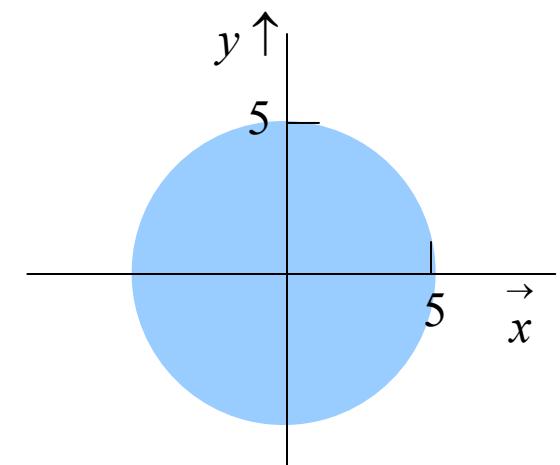
$$A = \{X + Y \leq 10\}$$

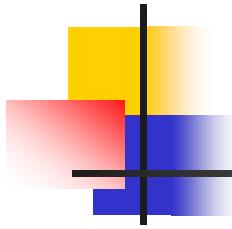


$$B = \{\min(X, Y) \leq 5\}$$



$$C = \{X^2 + Y^2 \leq 25\}$$





Independence

- Intuitively, if r.v.'s X and Y are “independent,” then events that involve only X should be independent of events that involve only Y . In other words, if A_1 is any event that involves X only and A_2 is any event that involves Y only, then

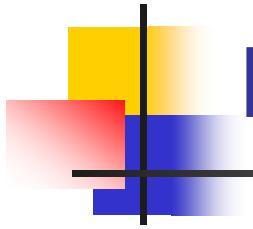
$$P[X \in A_1, Y \in A_2] = P[X \in A_1]P[Y \in A_2]$$

- In general, n r.v.'s are independent if

$$P[X_1 \in A_1, \dots, X_n \in A_n] = P[X_1 \in A_1] \cdots P[X_n \in A_n]$$

where A_k is an event that involves X_k only

if r.v.'s are independent, knowing the probabilities of the r.v.'s in isolation suffices to specify probabilities of joint events



Pairs of discrete r.v.'s vector r.v. $\mathbf{X} = (X, Y)$

$$S = \{(x_j, y_k), j = 1, 2, \dots, k = 1, 2, \dots\}$$

joint probability mass function:

$$\begin{aligned} p_{X,Y}(x_j, y_k) &= P[\{X = x_j\} \cap \{Y = y_k\}] \\ &\triangleq P[X = x_j, Y = y_k] \quad j = 1, 2, \dots \quad k = 1, 2, \dots \end{aligned}$$

gives probability of occurrence of pairs (x_j, y_k)

For any event A : $P[\mathbf{X} \in A] = \sum_{(x_j, y_k) \in A} \sum p_{X,Y}(x_j, y_k)$

$$\sum_{j=1}^{\infty} \sum_{k=1}^{\infty} p_{X,Y}(x_j, y_k) = P[S] = 1$$

Marginal PMF's

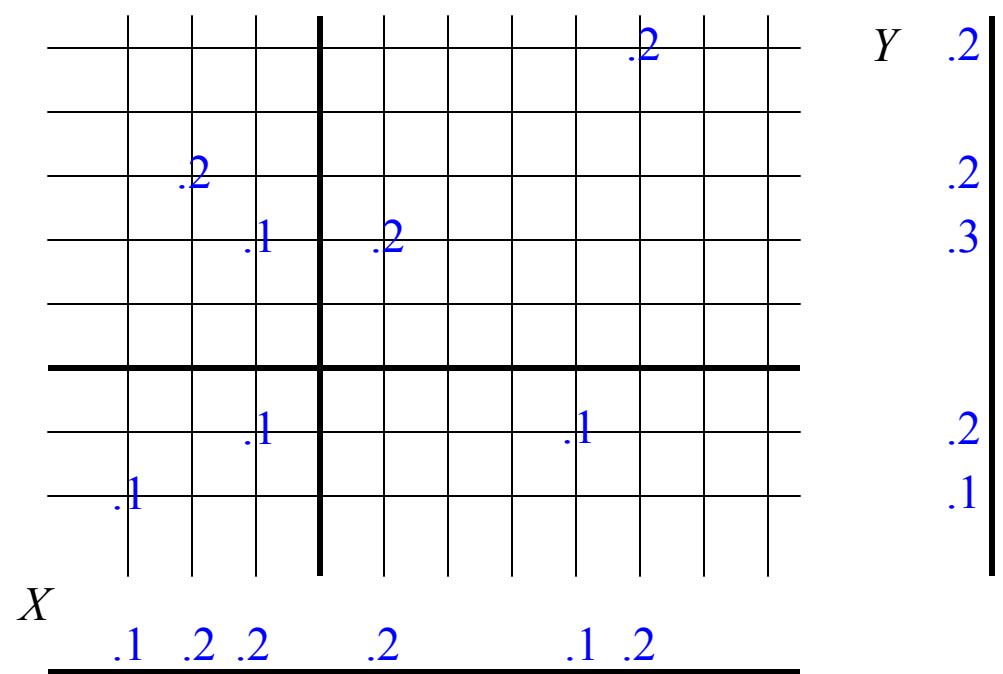
$$p_X(x_j) = p_{X,Y}(x_j, \text{any } y_k)$$

$$= P\left[\left\{\{X = x_j\} \cap \{Y = y_1\}\right\} \cup \left\{\{X = x_j\} \cap \{Y = y_2\}\right\} \cup \dots\right]$$

$$= \sum_{k=1}^{\infty} p_{X,Y}(x_j, y_k)$$

$$p_Y(y_k) = \sum_{j=1}^{\infty} p_{X,Y}(x_j, y_k)$$

marginal PMF's are insufficient – in general – for specifying joint PMF



Ex 4.6 tossing loaded dice

k joint PMF in units of $\frac{1}{42}$

j	1	2	3	4	5	6
1	2	1	1	1	1	1
2	1	2	1	1	1	1
3	1	1	2	1	1	1
4	1	1	1	2	1	1
5	1	1	1	1	2	1
6	1	1	1	1	1	2

7 7 7 7 7 7

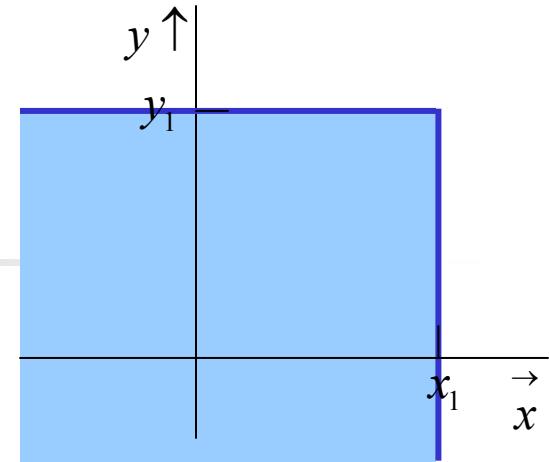
1 1 1 1 1 1

A graph showing two curves, one black and one blue, both increasing from left to right. The black curve is labeled '42' near its middle. The blue curve is labeled '6' near its end. A red arrow points downwards towards the blue curve.

marginals give no clue

Joint CDF of X and Y

- Basic building block:

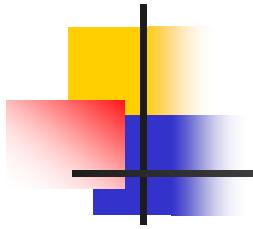


semi-infinite rectangle: $\{(x, y) : \{x \leq x_1\} \cap \{y \leq y_1\}\}$

joint cumulative distribution function of X and Y :

$$\begin{aligned} F_{X,Y}(x_1, y_1) &= P[\{x \leq x_1\} \cap \{y \leq y_1\}] && \text{product-form event} \\ &= P[X \leq x_1, Y \leq y_1] \end{aligned}$$

long-term proportion of times in which the outcome (x, y) falls in the blue rectangle, or the probability mass in it



properties of the joint CDF

i. $F_{X,Y}(x_1, y_1) \leq F_{X,Y}(x_2, y_2)$ $x_1 \leq x_2, y_1 \leq y_2$
non-decreasing in NE direction

ii. $F_{X,Y}(-\infty, y_1) = F_{X,Y}(x_2, -\infty) = 0$
neither X nor Y takes on values $< -\infty$

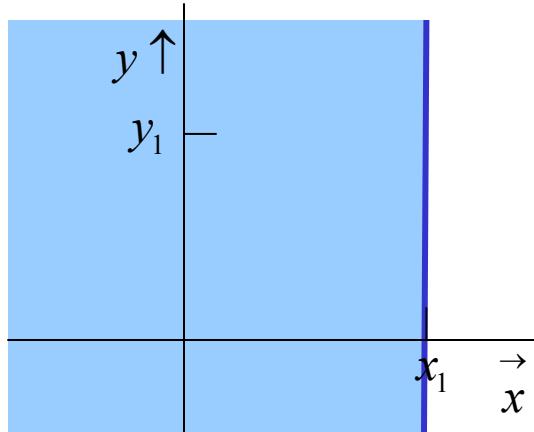
iii. $F_{X,Y}(\infty, \infty) = 1$
 X and Y can take on only values $< \infty$

iv. $F_X(x) = F_{X,Y}(x, \infty) = P[X \leq x, Y < \infty] = P[X \leq x]$
 $F_Y(y) = F_{X,Y}(\infty, y) = P[X < \infty, Y \leq y] = P[Y \leq y]$
marginal cumulative distribution functions

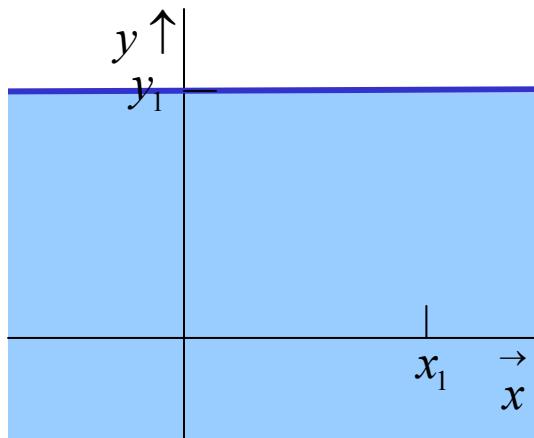
v. $\lim_{x \rightarrow a^+} F_{X,Y}(x, y) = F_{X,Y}(a, y)$ **continuous from N & E**
 $\lim_{y \rightarrow b^+} F_{X,Y}(x, y) = F_{X,Y}(x, b)$

Ex 4.8 marginals for continuous joint r.v.'s

Given $F_{X,Y}(x, y) = (1 - e^{-\alpha x})(1 - e^{-\beta y})u(x)u(y)$

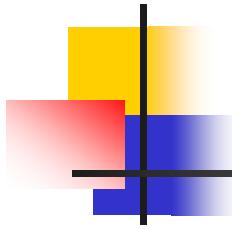


$$F_X(x) = \lim_{y \rightarrow \infty} F_{X,Y}(x, y) = (1 - e^{-\alpha x})u(x)$$



$$F_Y(y) = \lim_{x \rightarrow \infty} F_{X,Y}(x, y) = (1 - e^{-\beta y})u(y)$$

marginal CDFs are exponential distributions with parameters α and β respectively



Joint CDF in terms of joint PDF & v.v.

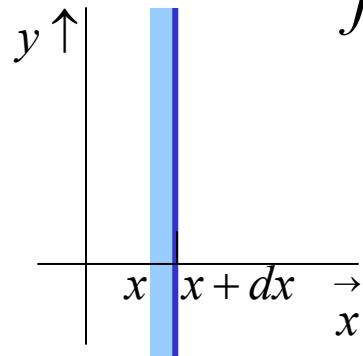
$$F_{X,Y}(x,y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(x',y') dx' dy'$$

↓
if jointly continuous r.v.'s

$$f_{X,Y}(x,y) = \frac{\partial^2 F_{X,Y}(x,y)}{\partial x \partial y}$$

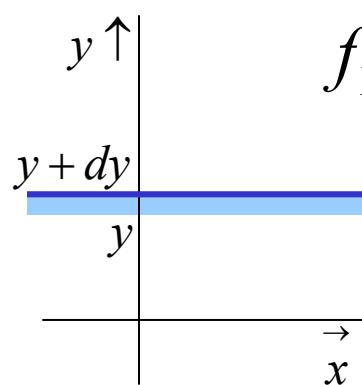
if the CDF is discontinuous
- or the partial derivatives are discontinuous -
then the joint PDF does not exist

Marginal PDF's from joint PDF



$$f_X(x) = \frac{d}{dx} F_{X,Y}(x, \infty) = \frac{d}{dx} \int_{-\infty}^x \left\{ \int_{-\infty}^{\infty} f_{X,Y}(x', y') dy' \right\} dx'$$

$$= \int_{-\infty}^{\infty} f_{X,Y}(x, y') dy' \quad \text{concentrating mass on x-axis}$$



$$f_Y(y) = \frac{d}{dy} F_{X,Y}(\infty, y) = \int_{-\infty}^{\infty} f_{X,Y}(x', y) dx' \quad \text{concentrating mass on y-axis}$$

marginal PDF's are obtained by integrating out the other variable(s)

$$f_{X,Y}(x,y) = \begin{cases} 1 & (x,y) \in [0,1] \times [0,1] \\ 0 & \text{o.w.} \end{cases}$$

Ex 4.10 CDF for jointly uniform r.v.'s

$$F_{X,Y}(x,y) = \int_{-\infty}^x \int_{-\infty}^y f_{X,Y}(x',y') dx' dy'$$

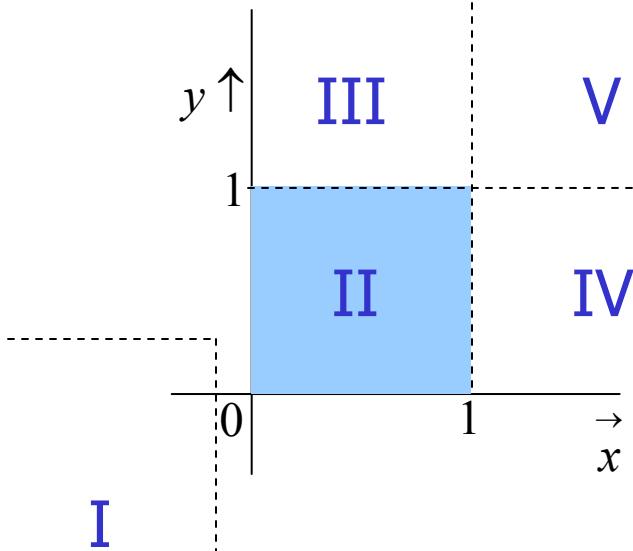
$$= 0 \quad \text{for } (x,y) \in I$$

$$= \int_0^x \int_0^y 1 dx' dy' = xy \quad \text{for } (x,y) \in II$$

$$= \int_0^x \int_0^1 1 dx' dy' = x \quad \text{for } (x,y) \in III$$

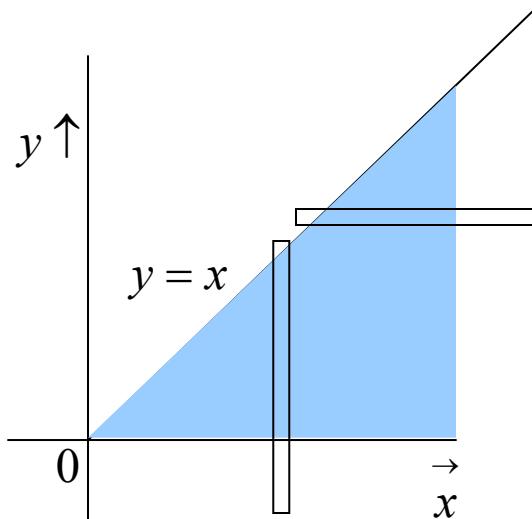
$$= \int_0^1 \int_0^y 1 dx' dy' = y \quad \text{for } (x,y) \in IV$$

$$= \int_0^1 \int_0^1 1 dx' dy' = 1 \quad \text{for } (x,y) \in V$$



Ex 4.11 finding marginal PDF's

$$f_{X,Y}(x,y) = \begin{cases} ce^{-x}e^{-y} & 0 \leq y \leq x < \infty \\ 0 & \text{o.w.} \end{cases}$$



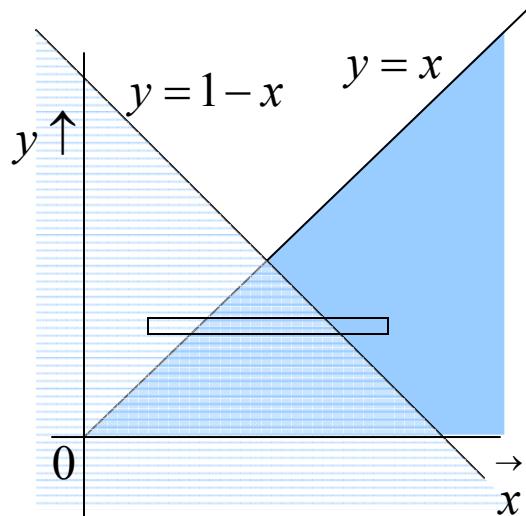
$$\begin{aligned} 1 &= \int_0^\infty \int_0^x ce^{-x}e^{-y} dy dx = \int_0^\infty ce^{-x} \int_0^x e^{-y} dy dx \\ &= \int_0^\infty ce^{-x} \frac{e^{-x} - 1}{-1} dx = \frac{c}{2} \Rightarrow c = 2 \end{aligned}$$

$$f_X(x) = \int_0^x 2e^{-x}e^{-y} dy = 2e^{-x} (1 - e^{-x}) \quad 0 \leq x < \infty$$

$$f_Y(y) = \int_y^\infty 2e^{-x}e^{-y} dx = 2e^{-y} (e^{-y}) = 2e^{-2y} \quad 0 \leq y < \infty$$

are these PDF's?

Ex 4.12 Find $P[X+Y \leq 1]$ in EX 4.11



$$\begin{aligned} P[X + Y \leq 1] &= \int_0^{0.5} \int_y^{1-y} 2e^{-x} e^{-y} dx dy \\ &= \int_0^{0.5} 2e^{-y} \int_y^{1-y} e^{-x} dx dy \\ &= \int_0^{0.5} 2e^{-y} \left[e^{-y} - e^{-(1-y)} \right] dy \\ &= -e^{-1} + 1 - e^{-1} = 1 - 2e^{-1} \end{aligned}$$

Ex 4.13 Find marginal PDFs of jointly Gaussian PDF

$$f_{X,Y}(x,y) = \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{(x^2 - 2\rho xy + y^2)}{2(1-\rho^2)}}$$

$$f_X(x) = \frac{e^{-\frac{x^2}{2(1-\rho^2)}}}{2\pi\sqrt{1-\rho^2}} \int_{-\infty}^{\infty} e^{-\frac{(-2\rho xy + y^2)}{2(1-\rho^2)}} dy$$

completing the square

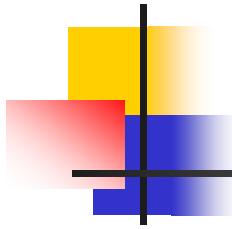
$$= \frac{e^{-\frac{-x^2(1-\rho^2)}{2(1-\rho^2)}}}{2\pi\sqrt{1-\rho^2}} \int_{-\infty}^{\infty} e^{-\frac{(\rho^2 x^2 - 2\rho xy + y^2)}{2(1-\rho^2)}} dy$$

$$= \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{1-\rho^2}\sqrt{2\pi}} e^{-\frac{(y-\rho x)^2}{2(1-\rho^2)}} dy = \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} = N(0,1)$$

symmetry

$$\downarrow$$

$$f_Y(y) = f_X(x)$$



Independence of two r.v.'s

discrete r.v.'s X and Y are independent

\Updownarrow iff

joint PMF is product of marginal PMFs

$$\forall x_j, y_k$$

Ex 4.15 another look at Ex 4.6 loaded dice

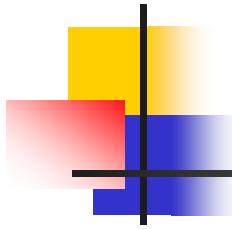
joint PMF in units of $\frac{1}{42}$

$j \backslash k$	1	2	3	4	5	6
1	2	1	1	1	1	1
2	1	2	1	1	1	1
3	1	1	2	1	1	1
4	1	1	1	2	1	1
5	1	1	1	1	2	1
6	1	1	1	1	1	2

marginals give no clue

not independent

$$p_{X,Y}(x_j, y_k) \neq p_X(x_j)p_Y(y_k) \quad \forall x_j, y_k$$



Independence of two r.v.'s in general

↑
↓
IFF

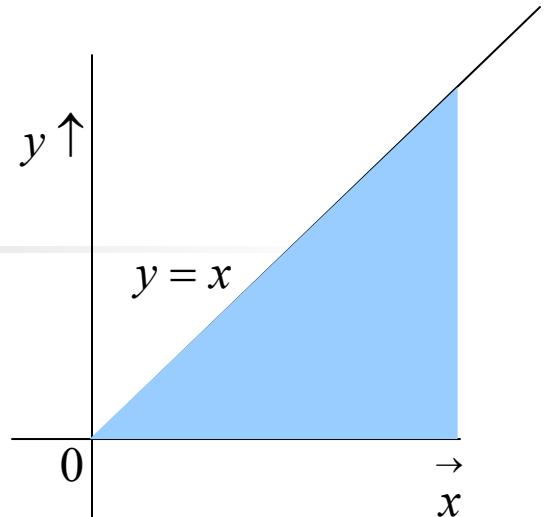
$$p_{X,Y}(x_j, y_k) = p_X(x_j)p_Y(y_k) \quad \forall x_j, y_k \quad \text{discrete}$$

$$F_{X,Y}(x, y) = F_X(x)F_Y(y) \quad \forall x, y$$

$$f_{X,Y}(x, y) = f_X(x)f_Y(y) \quad \forall x, y \quad \text{jointly continuous}$$

Ex 4.17 back to Ex 4.11

joint $f_{X,Y}(x,y) = \begin{cases} 2e^{-x}e^{-y} & 0 \leq y \leq x < \infty \\ 0 & o.w. \end{cases}$



$$f_X(x) = \int_0^x 2e^{-x}e^{-y} dy = 2e^{-x}(1 - e^{-x}) \quad 0 \leq x < \infty$$

marginal PDFs

$$f_Y(y) = \int_y^\infty 2e^{-x}e^{-y} dx = 2e^{-y}(e^{-y}) = 2e^{-2y} \quad 0 \leq y < \infty$$

product? $2e^{-x}e^{-y} \neq 2e^{-x}(1 - e^{-x})2e^{-2y}$

X and Y are not independent

Ex 4.18 back to Ex 4.13

$$\left. \begin{aligned}
 f_{X,Y}(x,y) &= \frac{1}{2\pi\sqrt{1-\rho^2}} e^{-\frac{(x^2 - 2\rho xy + y^2)}{2(1-\rho^2)}} \\
 f_X(x) &= \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} & f_Y(y) &= \frac{e^{-\frac{y^2}{2}}}{\sqrt{2\pi}} \\
 f_X(x)f_Y(y) &= \frac{e^{-\frac{x^2}{2}}}{\sqrt{2\pi}} \frac{e^{-\frac{y^2}{2}}}{\sqrt{2\pi}} = \frac{e^{-\frac{(x^2+y^2)}{2}}}{2\pi}
 \end{aligned} \right\} \equiv \text{ iff } \rho = 0$$

Conditional probability

$$P[Y \in A | X = x] = \frac{P[Y \in A, X = x]}{P[X = x]}$$

X discrete

$$F_Y(y | x_k) = \frac{P[Y \leq y, X = x_k]}{P[X = x_k]} \text{ for } P[X = x_k] > 0$$

if derivative exists

$$f_Y(y | x_k) = \frac{d}{dy} F_Y(y | x_k)$$

$$P[Y \in A | X = x_k] = \int_{y \in A} f_Y(y | x_k) dy$$

independence

$$F_Y(y | x) = F_Y(y)$$

$$f_Y(y | x) = f_Y(y)$$

Conditional probability

$$P[Y \in A | X = x] = \frac{P[Y \in A, X = x]}{P[X = x]}$$

X and Y discrete

$$p_Y(y_j | x_k) = P[Y = y_j | X = x_k] = \frac{P[X = x_k, Y = y_j]}{P[X = x_k]}$$

δ functions
~ PMF weights

$$= \begin{cases} \frac{p_{X,Y}(x_k, y_j)}{p_X(x_k)} & \text{for } P[X = x_k] = p_X(x_k) > 0 \\ 0 & \text{for } P[X = x_k] = p_X(x_k) = 0 \end{cases}$$

in general

$$P[Y \in A | X = x_k] = \sum_{y_j \in A} p_Y(y_j | x_k) = \sum_{y_j \in A} p_Y(y_j)$$

if also independent

Ex 4.20 back to Ex 4.14

- X is the input and Y the output of a communication channel. $P[Y < 0 | X = +1]?$

$$f_Y(y|1) \sim U[-1, 3] \rightarrow P[Y < 0 | X = +1] = \int_{-1}^0 \frac{1}{4} dy = \frac{1}{4}$$

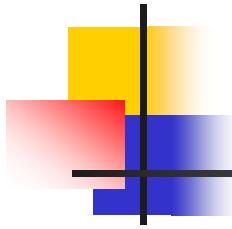
integrate over event

Conditional CDF of Y given $X=x$

$$F_Y(y|x) \triangleq \lim_{h \rightarrow 0} F_Y(y|x < X \leq x+h)$$

$$= \frac{P[Y \leq y, x < X \leq x+h]}{P[x < X \leq x+h]} = \frac{\int_{-\infty}^y \int_x^{x+h} f_{X,Y}(x', y') dx' dy'}{\int_x^{x+h} f_X(x') dx'}$$

$$\begin{aligned} & \simeq \frac{h \int_{-\infty}^y f_{X,Y}(x, y') dy'}{h f_X(x)} = \frac{\int_{-\infty}^y f_{X,Y}(x, y') dy'}{f_X(x)} \\ & = \int_{-\infty}^y f_Y(y') dy' = F_Y(y) \end{aligned} \quad \text{if independent}$$



Conditional PDF of Y given $X=x$

$$F_Y(y|x) = \frac{\int_{-\infty}^y f_{X,Y}(x,y') dy'}{f_X(x)}$$

$$f_Y(y|x) = \frac{d}{dy} F_Y(y|x) = \frac{f_{X,Y}(x,y)}{f_X(x)}$$

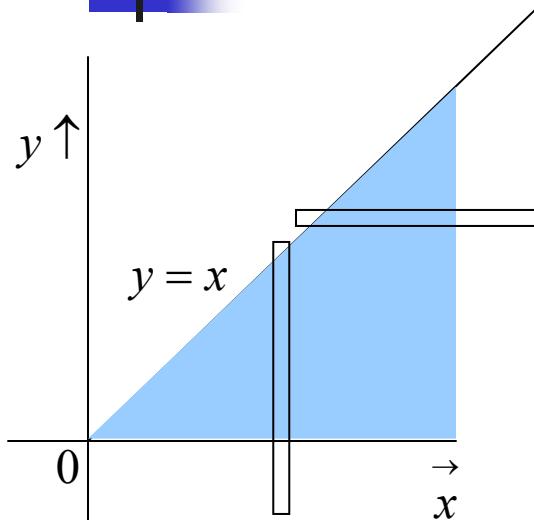
almost Bayes' rule

$$f_Y(y|x) dy = \frac{f_{X,Y}(x,y) dx dy}{f_X(x) dx}$$

if independent

$$f_Y(y|x) = f_Y(y)$$

Ex 4.21 back to Ex 4.11



$$f_{X,Y}(x,y) = \begin{cases} ce^{-x}e^{-y} & 0 \leq y \leq x < \infty \\ 0 & \text{o.w.} \end{cases}$$

$$f_X(x) = 2e^{-x} (1 - e^{-x}) \quad 0 \leq x < \infty$$

$$f_Y(y) = 2e^{-2y} \quad 0 \leq y < \infty$$

$$f_X(x | y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} = \frac{2e^{-x}e^{-y}}{2e^{-2y}} = e^{-(x-y)} ; y \leq x < \infty$$

$$f_Y(y | x) = \frac{f_{X,Y}(x,y)}{f_X(x)} = \frac{2e^{-x}e^{-y}}{2e^{-x}(1 - e^{-x})} = \frac{e^{-y}}{1 - e^{-x}} ; 0 \leq y < x$$