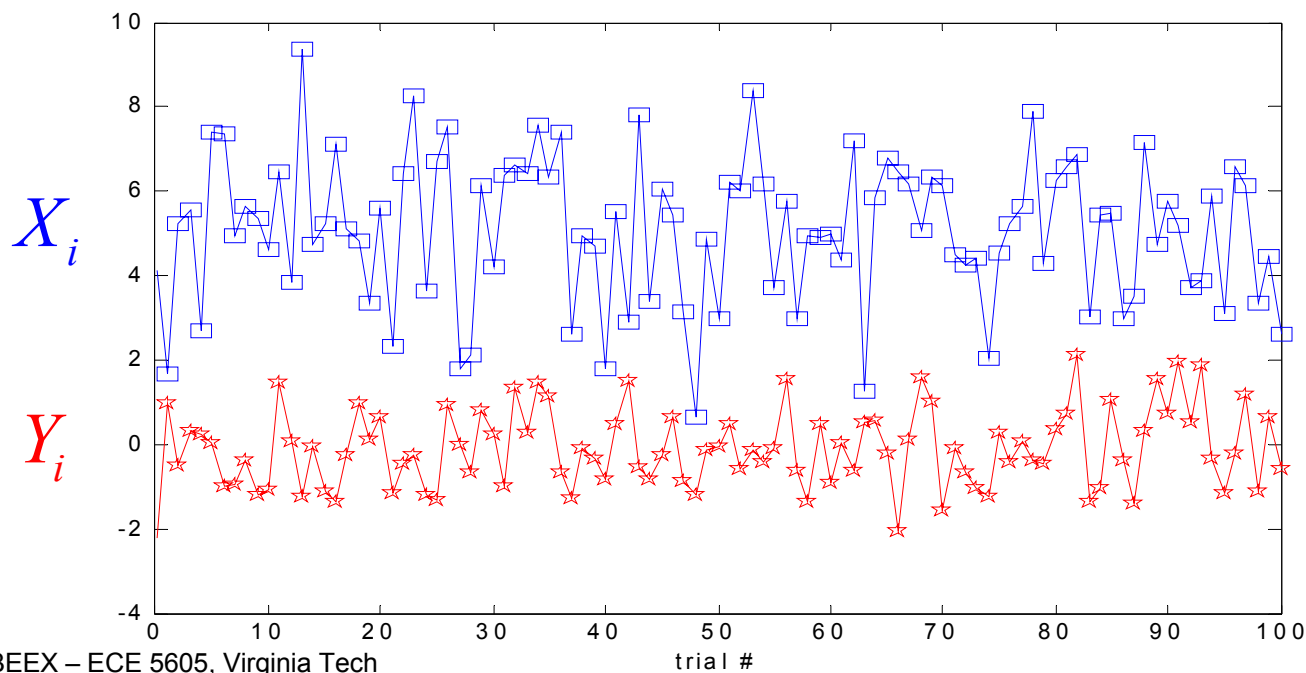


# Expected value of r.v.'s

- CDF or PDF are complete (probabilistic) descriptions of the behavior of a random variable. Sometimes we are interested in less information; in a partial characterization.



different center  
different spread



# Expected value or MEAN

$$E[X] \triangleq \int_{-\infty}^{\infty} t f_X(t) dt$$

$$E[X] \triangleq \sum_k x_k p_X(x_k)$$

defined if integral or sum converges absolutely

$$E[|X|] = \int_{-\infty}^{\infty} |t| f_X(t) dt < \infty$$

$$E[|X|] = \sum_k |x_k| p_X(x_k) < \infty$$

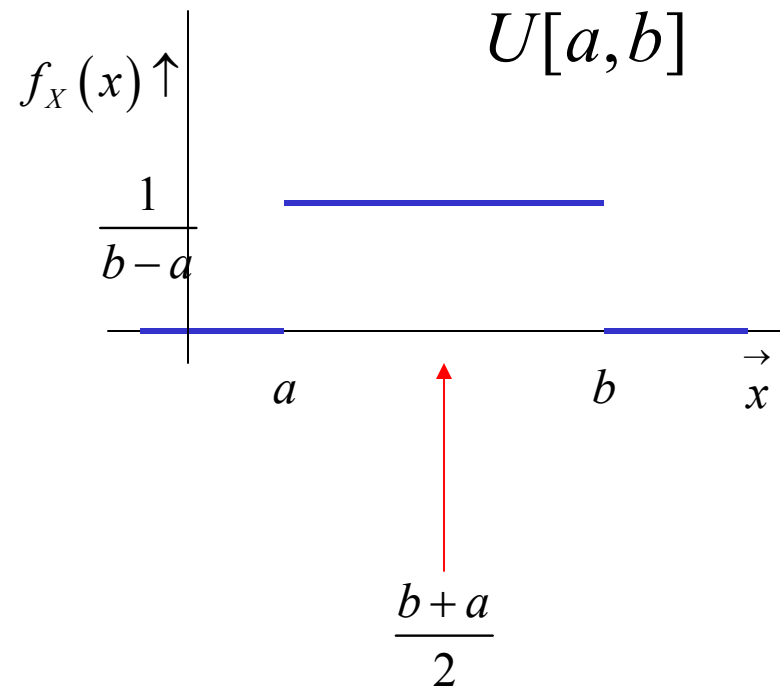
there are random variables for which the above do not converge  
we then say "the mean does not exist"

$E[X]$  represents the "center of mass"

arithmetic average of large # independent observations of a r.v.  
will tend to the mean; it's like the "average of  $X$ "

## Ex 3.29 mean of $U[a,b]$

$$\begin{aligned} E[X] &\triangleq \int_{-\infty}^{\infty} t f_X(t) dt \\ &= \int_a^b t \frac{1}{b-a} dt \\ &= \frac{b^2 - a^2}{2(b-a)} = \frac{b+a}{2} \end{aligned}$$



midpoint of interval  $[a,b]$



## Ex 3.31 mean of exponential r.v.

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- Time  $X$  between customer arrivals at a service station has an exponential PDF with parameter  $\lambda$ . Find the mean inter-arrival time.

$$E[X] = \int_0^{\infty} \underbrace{t \lambda e^{-\lambda t}}_u dv = t e^{-\lambda t} \Big|_0^{\infty} + \int_0^{\infty} e^{-\lambda t} dt$$
$$= \lim_{t \rightarrow \infty} t e^{-\lambda t} - 0 + \frac{e^{-\lambda t}}{-\lambda} \Big|_0^{\infty} = \frac{1}{\lambda}$$

$\int u dv = uv - \int v du$

As  $\lambda$  is customer arrival rate in customers per second, the mean inter-arrival time of  $\lambda^{-1}$  seconds/customer makes sense!

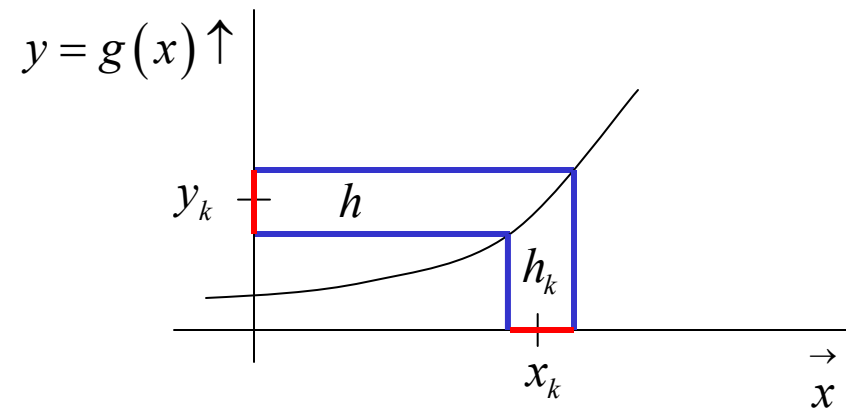
# Expected value of a function of a r.v.

$$E[Y] \triangleq \int_{-\infty}^{\infty} t f_Y(t) dt$$

alternative in terms of  $X$

$$\begin{aligned} E[Y] &\approx \sum_k y_k f_Y(y_k) h \\ &= \sum_k g(x_k) f_X(x_k) h_k \\ &= \int_{-\infty}^{\infty} g(x) f_X(x) dx \end{aligned}$$

limit  $h \rightarrow 0$



$$\Theta \sim U(0, 2\pi]$$

$$Y = a \cos(\omega t + \Theta)$$



## Ex 3.33

- Sampling a sinusoid with random phase. Find the expected value of  $Y$  and of  $Y^2$ , the power of  $Y$ .

$$\begin{aligned} E[Y] &= E[a \cos(\omega t + \Theta)] \\ &= \int_0^{2\pi} a \cos(\omega t + \theta) \frac{1}{2\pi} d\theta \\ &= \frac{-a}{2\pi} \sin(\omega t + \theta) \Big|_0^{2\pi} \\ &= \frac{-a}{2\pi} [\sin(\omega t + 2\pi) - \sin(\omega t)] = 0 \end{aligned}$$

$$\begin{aligned} E[Y^2] &= E[a^2 \cos^2(\omega t + \Theta)] \\ &= \frac{a^2}{2} E[1 + \cos(2\omega t + 2\Theta)] \\ &= \frac{a^2}{2} \left[ 1 + \int_0^{2\pi} \cos(2\omega t + 2\theta) \frac{1}{2\pi} d\theta \right] \\ &= \frac{a^2}{2} \end{aligned}$$

agreement with time-averages: "DC" value of 0, power  $a^2/2$



## Variance of $X$

- Provides information about a random variable – in addition to its mean value – regarding its deviations from the mean

$$D \triangleq X - E[X] \quad \text{deviation from the mean}$$

$$VAR[X] \triangleq E\left[\left(X - E[X]\right)^2\right]$$

$$STD[X] \triangleq \sqrt{VAR[X]}$$

measures of “width/spread”

variance of r.v.

standard deviation of r.v.

$$\begin{aligned} VAR[X] &= E\left[X^2 - 2E[X]X + E^2[X]\right] \\ &= E\left[X^2\right] - 2E[X]E[X] + E^2[X] \\ &= E\left[X^2\right] - E^2[X] \end{aligned}$$



## Ex 3.36 variance of $X \sim U[a, b]$

$$E[X] = \frac{b+a}{2}$$

$$VAR[X] = \frac{1}{b-a} \int_a^b \left( x - \frac{b+a}{2} \right)^2 dx$$

$$= \frac{1}{b-a} \int_{-(b-a)/2}^{(b-a)/2} y^2 dy = \frac{y^3}{3(b-a)} \Big|_{-(b-a)/2}^{(b-a)/2} = \frac{(b-a)^2}{12}$$

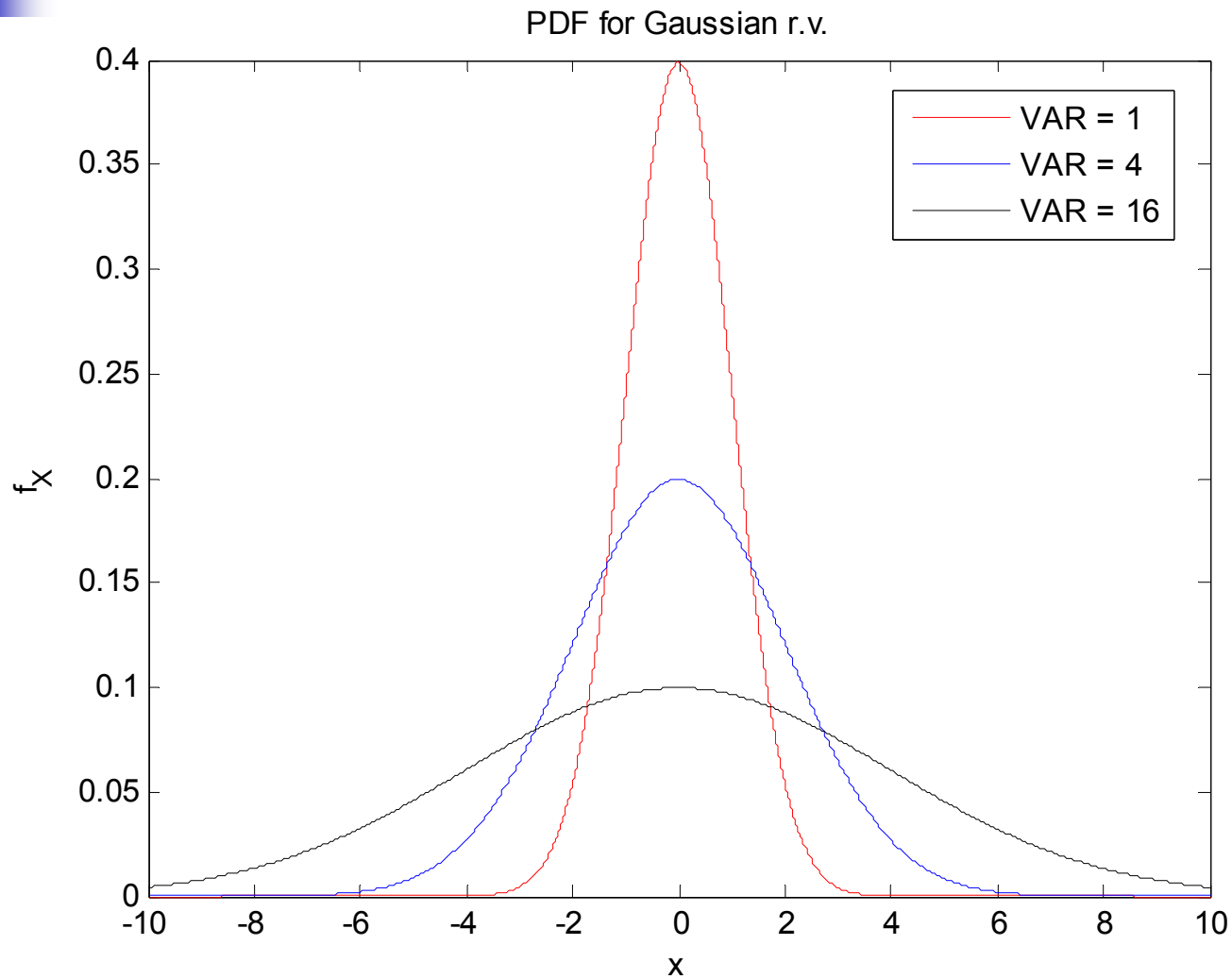
$$U[-2, 4] \rightarrow E[X] = \frac{4+(-2)}{2} = 1; VAR[X] = \frac{(4-(-2))^2}{12} = 3$$



## 3.38 variance of Gaussian r.v.

$$\begin{aligned} f_X(x) &= \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-m)^2}{2\sigma^2}} \quad -\infty < x < \infty \\ \int_{-\infty}^{\infty} f_X(x) dx &= \int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-m)^2}{2\sigma^2}} dx \\ \sigma\sqrt{2\pi} &= \int_{-\infty}^{\infty} e^{-\frac{(x-m)^2}{2\sigma^2}} dx \\ \frac{d}{d\sigma} \sigma\sqrt{2\pi} &= \frac{d}{d\sigma} \int_{-\infty}^{\infty} e^{-\frac{(x-m)^2}{2\sigma^2}} dx \\ \sqrt{2\pi} &= \int_{-\infty}^{\infty} \frac{(x-m)^2}{\sigma^3} e^{-\frac{(x-m)^2}{2\sigma^2}} dx \\ \text{VAR}[X] &= \frac{1}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} (x-m)^2 e^{-\frac{(x-m)^2}{2\sigma^2}} dx = \sigma^2 \end{aligned}$$

# Effect of VAR on the PDF of a Gaussian r.v.





# Mean & variance

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- Are the two most important parameters for (partially) characterizing the PDF of a r.v.

- Others are sometimes used, e.g.

$$\text{skewness} \triangleq \frac{E\left[\left(X - E[X]\right)^3\right]}{\sigma^3}$$

which measures the degree of asymmetry about the mean

- *Skewness is zero for a symmetric PDF*

- Each involves  $n^{\text{th}}$  moment of r.v.  $X$ :  $E[X^n] = \int_{-\infty}^{\infty} x^n f_X(x) dx$

- $VAR[X] = E[X^2] - E^2[X]$

- Under certain conditions, a PDF is completely specified if all moments are known (more on that later)

$$VAR[c] = 0 \quad VAR[X + c] = VAR[X] \quad VAR[cX] = c^2 VAR[X]$$