

Random or stochastic processes

Definition

- Random experiment specified by $\zeta \in S$, by the events defined on S , and by the probabilities on these events. To each $\zeta \in S$, a function of time is assigned according to some rule.

$$X(t, \zeta) \quad t \in I$$

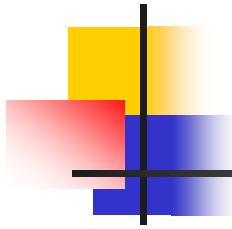
$X(t, \zeta)|_{\zeta}$ is a function of time t
realization, sample path, sample function

$X(t, \zeta)|_{t=t_k}$ is a function of ζ , i.e. an r.v.

$\{X(t, \zeta), t \in I\}$ an indexed family of r.v.'s
or random process
or stochastic process

if I is countable: discrete time SP

if I is continuous (non-negative) real line: continuous SP



Ex 6.1 discrete-time SP

$$\zeta \in S = (0, 1]$$

$$\forall \zeta \in S \rightarrow \zeta = \sum_{n=1}^{\infty} b_n 2^{-n} \quad \text{binary expansion}$$

$$\text{discrete-time SP: } X(n, \zeta) = b_n ; n = 1, 2, \dots$$

Ex 6.2 continuous-time SP

$$\zeta \in S = [-1, 1]$$

$$X(t, \zeta) = \zeta \cos(2\pi t); -\infty < t < \infty$$

sinusoids with amplitude ζ

$$\zeta \in S = [-\pi, \pi]$$

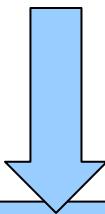
$$Y(t, \zeta) = \cos(2\pi t + \zeta); -\infty < t < \infty$$

time-shifted versions of cos

Specifying a SP

sample the SP $X(t, \zeta)|_{t=t_i} \rightarrow$ r.v.'s X_1, X_2, \dots, X_k

joint behavior of the SP at these k time instants is specified by the joint CDF of the vector r.v. $\mathbf{X} = (X_1, X_2, \dots, X_k)$



we can now compute probabilities as before

a SP is specified by the collection of k^{th} -order joint CDFs:

$$F_{X_1, \dots, X_k}(x_1, \dots, x_k) = P[X_1 \leq x_1, X_2 \leq x_2, \dots, X_k \leq x_k]$$

for any k and any choice of sampling instants t_1, \dots, t_k

$$p_{X_1, \dots, X_k}(x_1, \dots, x_k) = P[X_1 = x_1, X_2 = x_2, \dots, X_k = x_k] \quad \text{for discrete-time SP}$$

$$f_{X_1, \dots, X_k}(x_1, \dots, x_k) \quad \text{for continuous-time SP}$$

partial characterizations using moments

Mean, autocorrelation, autocovariance

$$m_X(t) = E[X(t)] = \int_{-\infty}^{\infty} xf_{X(t)}(x)dx$$

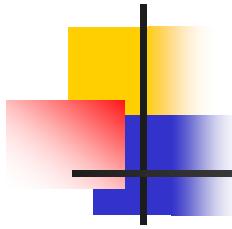
$$R_X(t_1, t_2) = E[X(t_1)X(t_2)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xyf_{X(t_1)X(t_2)}(x, y)dxdy$$

$$\begin{aligned} C_X(t_1, t_2) &= E[\{X(t_1) - m_X(t_1)\}\{X(t_2) - m_X(t_2)\}] \\ &= R_X(t_1, t_2) - m_X(t_1)m_X(t_2) \end{aligned}$$

$$VAR[X(t)] = E[\{X(t) - m_X(t)\}^2] = C_X(t, t)$$

$$\rho_X(t_1, t_2) = \frac{C_X(t_1, t_2)}{\sqrt{C_X(t_1, t_1)C_X(t_2, t_2)}} \implies |\rho_X(t_1, t_2)| \leq 1$$

~predicting one r.v. as a linear function from another³



Ex 6.6 random amplitude sinusoid

$$X(t) = A \cos 2\pi t$$

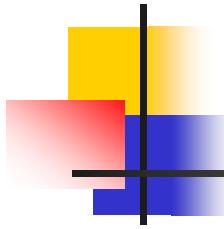
$$m_X(t) = E[A \cos 2\pi t] = E[A] \cos 2\pi t$$

mean varies with t

process is always =0 for t -values
for which $\cos 2\pi t=0$

$$\begin{aligned} R_X(t_1, t_2) &= E[A \cos(2\pi t_1) A \cos(2\pi t_2)] \\ &= E[A^2] \cos(2\pi t_1) \cos(2\pi t_2) \end{aligned}$$

$$\begin{aligned} C_X(t_1, t_2) &= R_X(t_1, t_2) - m_X(t_1) m_X(t_2) \\ &= \{E[A^2] - E[A]E[A]\} \cos(2\pi t_1) \cos(2\pi t_2) \\ &= VAR(A) \cos(2\pi t_1) \cos(2\pi t_2) \end{aligned}$$



Ex 6.7 random phase sinusoid

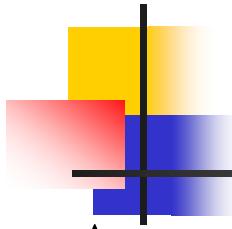
$$X(t) = \cos(\omega t + \Theta); \Theta \sim U(-\pi, \pi)$$

$$m_X(t) = E[\cos(\omega t + \Theta)] = \frac{1}{2\pi} \int_{-\pi}^{\pi} \cos(\omega t + \theta) d\theta = 0$$

mean is constant

$$\begin{aligned} C_X(t_1, t_2) &= R_X(t_1, t_2) = E[\cos(\omega t_1 + \Theta) \cos(\omega t_2 + \Theta)] \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \frac{1}{2} \left\{ \cos(\omega(t_1 - t_2)) + \cos(\omega(t_1 + t_2) + 2\theta) \right\} d\theta \\ &= \frac{1}{2} \cos(\omega(t_1 - t_2)) \end{aligned}$$

covariance depends on absolute time difference only



Gaussian SP

A random process $X(t)$ is a Gaussian SP

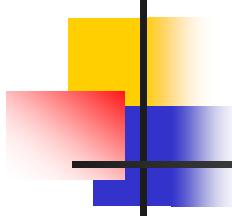
if for any k and any choice of sampling instants $t_1 < t_2 < \dots < t_k$

the r.v.'s $X_1 = X(t_1), X_2 = X(t_2), \dots, X_k = X(t_k)$

are jointly Gaussian r.v.'s

discrete & continuous

$$f_{\mathbf{X}}(\mathbf{x}) \triangleq f_{X_1, X_2, \dots, X_k}(x_1, x_2, \dots, x_k) \\ = \frac{\exp\left\{-\frac{1}{2}(\mathbf{x}-\mathbf{m})^T K^{-1}(\mathbf{x}-\mathbf{m})\right\}}{(2\pi)^{k/2} |K|^{1/2}}$$

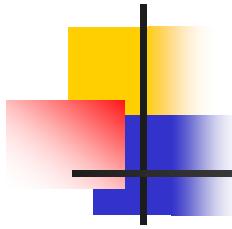


k jointly Gaussian r.v.'s

r.v.'s X_1, X_2, \dots, X_k are said to be jointly Gaussian if their joint PDF is given by

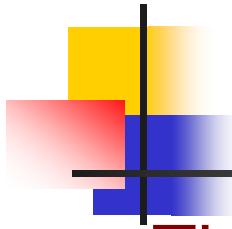
$$f_{\mathbf{X}}(\mathbf{x}) \triangleq f_{X_1, X_2, \dots, X_k}(x_1, x_2, \dots, x_k) = \frac{\exp\left\{-\frac{1}{2}(\mathbf{x} - \mathbf{m})^T K^{-1}(\mathbf{x} - \mathbf{m})\right\}}{(2\pi)^{k/2} |K|^{\frac{1}{2}}}$$

$$\mathbf{m} = \begin{bmatrix} m_X(t_1) \\ m_X(t_2) \\ \vdots \\ m_X(t_k) \end{bmatrix} \quad K = \left\{ C_X(t_i, t_j) \right\}_{i,j}$$



Multiple SPs

- Joint behavior of two or more SPs is specified by the collection of joint distributions for all possible choices of time samples of the SPs
- e.g. for $X(t)$ and $Y(t)$ specify all possible joint density functions of $X(t_1), \dots, X(t_k)$ and $Y(t'_1), \dots, Y(t'_j)$ for all k, j and all choices of t_1, \dots, t_k and t'_1, \dots, t'_j



Multiple SPs

- The SPs $X(t)$ and $Y(t)$ are said to be independent if the vector r.v.'s $(X(t_1), \dots, X(t_k))$ and $(Y(t'_1), \dots, Y(t'_j))$ are independent for all k, j and all choices of t_1, \dots, t_k and t'_1, \dots, t'_j

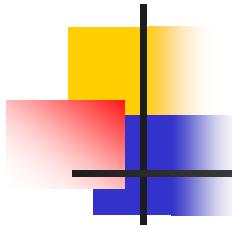
cross-correlation $R_{XY}(t_1, t_2) = E[X(t_1)Y(t_2)]$

SPs $X(t)$ and $Y(t)$ are said to be **orthogonal** if $R_{XY}(t_1, t_2) = 0 \forall t_1, t_2$

cross-covariance

$$\begin{aligned} C_{XY}(t_1, t_2) &= E[\{X(t_1) - m_X(t_1)\}\{Y(t_2) - m_Y(t_2)\}] \\ &= R_{XY}(t_1, t_2) - m_X(t_1)m_Y(t_2) \end{aligned}$$

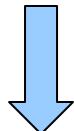
SPs $X(t)$ and $Y(t)$ are said to be **uncorrelated** if $C_{XY}(t_1, t_2) = 0 \forall t_1, t_2$



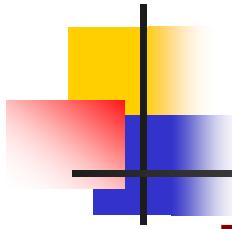
Ex 6.10 signal + noise

$$Y(t) = X(t) + N(t)$$

cross-correlation



$$\begin{aligned} R_{XY}(t_1, t_2) &= E[X(t_1)Y(t_2)] \\ &= E[X(t_1)\{X(t_2) + N(t_2)\}] \\ &= E[X(t_1)X(t_2)] + E[X(t_1)N(t_2)] \\ &= R_X(t_1, t_2) + E[X(t_1)]E[N(t_2)] \quad \text{independence} \\ &= R_X(t_1, t_2) + m_X(t_1)m_N(t_2) \end{aligned}$$

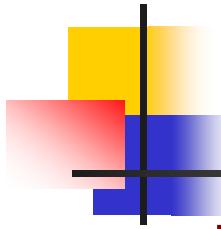


Stationary random processes

- Essentially...the nature of the randomness does not change with time
- Assuming the process started at $t=-\infty$...
A discrete-time or continuous-time SP $X(t)$ is stationary if the joint distribution of any set of samples does not depend on the placement of the time origin

$$F_{X(t_1), \dots, X(t_k)}(x_1, \dots, x_k) = F_{X(t_1+\tau), \dots, X(t_k+\tau)}(x_1, \dots, x_k); \forall \tau, \forall k, \forall t_1, \dots, t_k$$

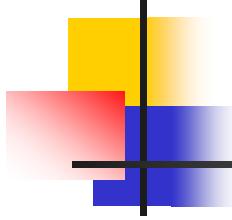
if SP starts at definite time, eg. $t=0$ or $n=0$, joint CDF does not change under time shifts to the right



Jointly stationary SPs

- Two processes $X(t)$ and $Y(t)$ are said to be jointly stationary if

$$\begin{aligned} F_{X(t_1), \dots, X(t_k), Y(t'_1), \dots, Y(t'_j)}(x_1, \dots, x_k, y_1, \dots, y_j) &= \\ &= F_{X(t_1+\tau), \dots, X(t_k+\tau), Y(t'_1+\tau), \dots, Y(t'_j+\tau)}(x_1, \dots, x_k, y_1, \dots, y_j) \\ &\quad \forall \tau, \forall k, j, \forall t_1, \dots, t_k, \forall t'_1, \dots, t'_k \end{aligned}$$



First-order CDF of stationary SP

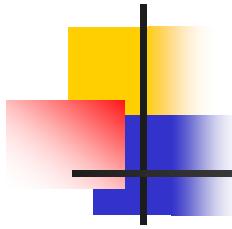
$$F_{X(t)}(x) = F_{X(t+\tau)}(x); \forall \tau, \forall t$$

$$= F_X(x)$$

$$\downarrow$$
$$m_X(t) = E[X(t)] = m; \forall t$$

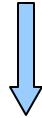
$$VAR(X(t)) = E[(X(t) - m)^2] = \sigma^2; \forall t$$

mean and variance are constant



Second-order CDF of stationary SP

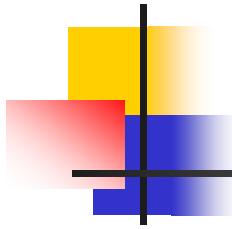
$$F_{X(t_1), X(t_2)}(x_1, x_2) = F_{X(t_1 + \tau), X(t_2 + \tau)}(x_1, x_2)$$
$$= F_{X(0), X(t_2 - t_1)}(x_1, x_2); \forall t_1, t_2$$



$$R_X(t_1, t_2) = R_X(t_2 - t_1); \forall t_1, t_2$$

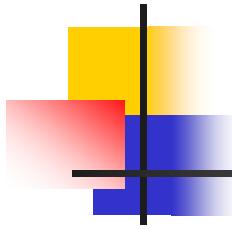
$$C_X(t_1, t_2) = C_X(t_2 - t_1); \forall t_1, t_2$$

autocorrelation and covariance depend only on time difference



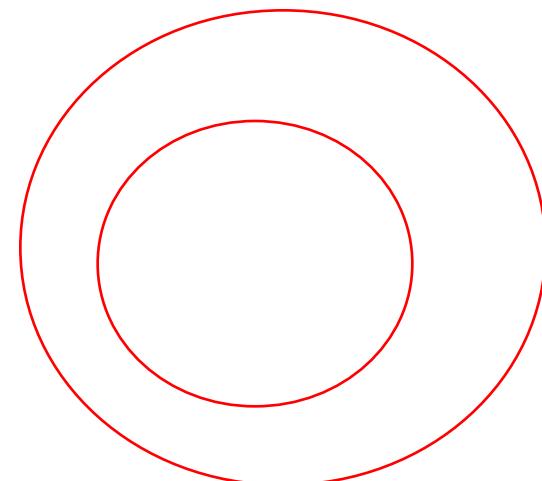
Wide-sense stationary (WSS) SP

- Often we cannot determine whether a process is stationary, but we can determine whether
 - *the mean is constant* $m_X(t) = m ; \forall t$
 - *the autocovariance (equivalently the autocorrelation) is a function of time difference only* $C_X(t_1, t_2) = C_X(t_1 - t_2); \forall t_1, t_2$
- SP is WSS if both conditions hold



Jointly WSS SPs

- $X(t)$ and $Y(t)$ are jointly WSS SPs if
 - $X(t)$ is WSS
 - $Y(t)$ is WSS
 - $C_{XY}(t_1, t_2) = C_{XY}(t_1 - t_2) = C_{XY}(\tau)$
- All stationary SPs are WSS



Autocorrelation function of WSS SP

- Plays a crucial role in the design of linear signal processing algorithms

$$i \quad R_X(0) = E[X^2(t)] \quad \text{average power of the process}$$

$$\text{ii} \quad R_X(\tau) = E[X(t+\tau)X(t)] = E[X(t)X(t+\tau)] = R_X(-\tau)$$

an even function of τ

$$\text{iii} \quad E^2[XY] \leq E[X^2]E[Y^2] \quad \forall X, Y$$

$$R_X^2(\tau) = E^2[X(t+\tau)X(t)] \leq E[X^2(t+\tau)]E[X^2(t)] = R_X^2(0)$$

\downarrow

$$|R_X(\tau)| \leq R_X(0)$$

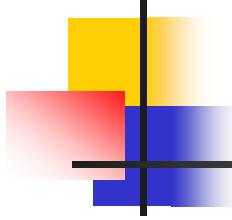
autocorrelation is max at $\tau=0$

- If $R_x(d)=R_x(0) \rightarrow R_x(\tau)$ is periodic with period equal to d .
- Let $X(t) = m + N(t)$

$N(t)$ is a zero-mean process $\Rightarrow R_N(\tau) \xrightarrow{\tau \rightarrow \infty} 0$

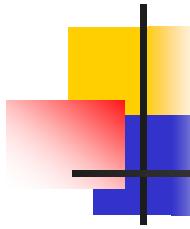
$$\begin{aligned} R_X(\tau) &= E\left[\{m + N(t + \tau)\}\{m + N(t)\}\right] \\ &= m^2 + 2mE[N(t)] + R_N(\tau) \\ &= m^2 + R_N(\tau) \xrightarrow{\tau \rightarrow \infty} m^2 \end{aligned}$$

$R_X(\tau)$ approaches m^2 (the square of the mean) as $\tau \rightarrow \infty$



Autocorrelation function of WSS SP

- Can have three types of component:
- A component that vanishes as $\tau \rightarrow \infty$
- A periodic component
- A component due to a nonzero mean



Ex 6.30

$$R_X(\tau) = e^{-2\alpha|\tau|}$$

$X(t)$ is zero-mean and $R_X(\tau) \xrightarrow{\tau \rightarrow \infty} 0$

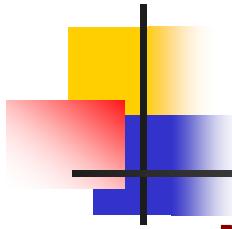
$$R_Y(\tau) = \frac{a^2}{2} \cos(2\pi f_0 \tau) \quad \text{random phase sinusoid}$$

$Y(t)$ is zero-mean and $R_X(\tau)$ has period f_0^{-1}

$$R_Z(\tau) = R_X(\tau) + R_Y(\tau) + m^2$$

$Z(t) = X(t) + Y(t) + m$; $X(t)$ and $Y(t)$ independent

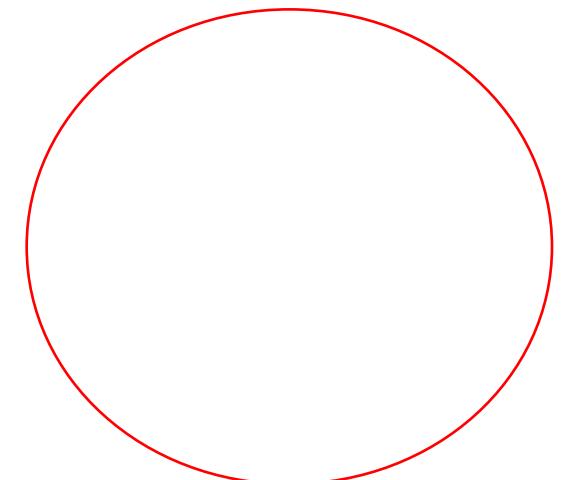
Note that it is not possible to determine which fraction of the mean is contributed by the individual components



WSS Gaussian SPs

- Mean is not a function of t
 - Autocovariance is a function of time difference only
- ↓
- Joint PDF depends on mean and autocovariance only
- ↓
- A WSS Gaussian SP is stationary

all we need is m and $C_X(\tau)$



Ex 6.31

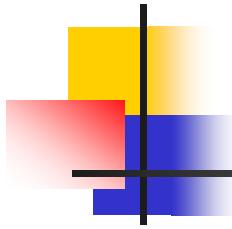
X_n is iid sequence of $N(0, \sigma^2)$ r.v.'s. Y_n is the average of two consecutive values

$$Y_n = \frac{X_n + X_{n-1}}{2} \implies m_Y = 0.5E[X_n + X_{n-1}] = 0$$

$$\begin{aligned} C_Y(i, j) &= E[Y_i Y_j] = 0.25E[(X_i + X_{i-1})(X_j + X_{j-1})] \\ &= 0.25\{E[X_i X_j] + E[X_i X_{j-1}] + E[X_{i-1} X_j] + E[X_{i-1} X_{j-1}]\} \\ &= \frac{1}{2}\sigma^2\delta_{i-j} + \frac{1}{4}\sigma^2\delta_{i-j+1} + \frac{1}{4}\sigma^2\delta_{i-j-1} \quad Y_n \text{ is WSS} \end{aligned}$$

Y_n is a linear trafo of Gaussian r.v.'s, so Y_n is Gaussian

joint PDF is Gaussian, specified by m_Y and $C_Y(i, j)$



Time averages & ergodicity

$$\hat{m}_X(t) = \frac{1}{N} \sum_{i=1}^N X(t, \zeta_i)$$

ensemble averaging
repeating the experiment
from many realizations

$$\langle X(t) \rangle_T = \frac{1}{2T} \int_{-T}^T X(t, \zeta) dt$$

time averaging
based on single realization



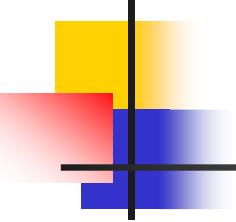
an ergodic theorem states conditions under which a time average converges when the observation interval becomes large

we're interested in ergodic theorems that state when time averages converge to the ensemble average or expected value

$$X(t) = A \cos(2\pi f_c t + \vartheta)$$

ϑ is U.RV from $0 \longrightarrow 2\pi$

Is $X(t)$ mean and autocorrelation ergodic?



Mean, Autocorrelation of sine wave with random phase

$$X(t) = A \cos(2\pi f_c t + \theta)$$

A and f_c are constants and θ is a RV that is uniformly distributed over range of 0 and 2π

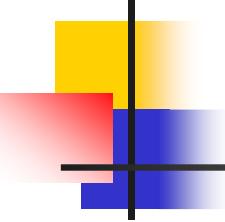
$$m_x(t) = \int_0^{2\pi} A \cos(2\pi f_c t + \theta) \frac{1}{2\pi} d\theta$$

$$m_x(t) = A [\sin(2\pi f_c t + \theta)]_0^{2\pi} = zero$$

$$\langle X(t)_T \rangle = \int_0^T A \cos(2\pi f_c t + \theta) dt$$

$$\langle X(t)_T \rangle = A \left[\frac{\sin(2\pi f_c t + \theta)}{2\pi f_c} \right]_0^T = zero$$

Ergodic



Mean, Autocorrelation of sine wave with random phase

$$X(t) = A \cos(2\pi f_c t + \theta)$$

A and f_c are constants and θ is a RV that is uniformly distributed over range of 0 and 2π

$$R_x(t_1, t_2) = E[X(t_1)X(t_2)] = A^2 E[\cos(2\pi f_c t_1 + \theta)\cos(2\pi f_c t_2 + \theta)]$$

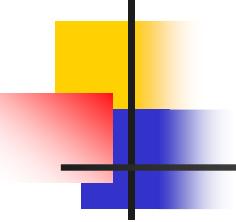
$$R_x(t_1, t_2) = \frac{A^2}{2} E[\cos(2\pi f_c(t_1 - t_2)) + \cos(2\pi f_c(t_1 + t_2) + 2\theta)]$$

$$R_x(t_1, t_2) = \frac{A^2}{2} \cos(2\pi f_c(t_1 - t_2)) = R_x(\tau) = \frac{A^2}{2} \cos(2\pi f_c(\tau)) \text{ W.S Stationary}$$

$$\langle X(t)X(t+\tau) \rangle = A^2 \frac{1}{2T} \int_{-T}^T \cos(2\pi f_c t + \theta) \cos(2\pi f_c(t+\tau) + \theta) dt$$

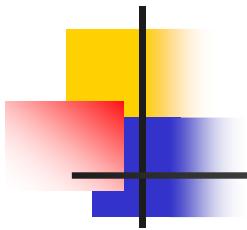
$$\langle X(t)X(t+\tau) \rangle = A^2 \frac{1}{2T} \int_{-T}^T [\cos(2\pi f_c \tau) + \cos(2\pi f_c(2t+\tau) + 2\theta)] dt$$

$$\langle X(t)X(t+\tau) \rangle = \frac{A^2}{2} \cos(2\pi f_c(\tau)) = R_x(\tau) \quad \text{Ergodic}$$



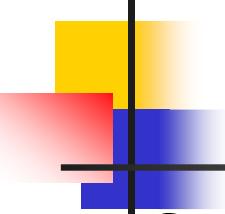
Properties of Gaussian RP

1. If a Gaussian process $X(t)$ is applied to a LTIS then the output is also a Gaussian Process.
2. Gaussian Random process defined at a set of time instants is completely defined by the vector mean M and the covariance matrix C
3. If $X(t)$ is a Gaussian WSS RP then it is a SS RP
4. If $x(t)$ is a Gaussian RP with uncorrelated RV then they are also independent.



Cross Correlation function

- Consider two random processes $X(t)$ and $Y(t)$ with autocorrelation functions $R_x(t_1, t_2)$, $R_y(t_1, t_2)$ respectively
- The cross correlation function of $X(t)$ and $Y(t)$ is defined by $R_{xy}(t_1, t_2) = E[X(t_1)Y(t_2)]$ and $R_{yx}(t_1, t_2) = E[Y(t_1)X(t_2)]$.
- If $X(t)$ and $Y(t)$ are WSS then $R_{xy}(t_1, t_2) = R_{xy}(\tau)$
- $R_{xy}(\tau) = R_{yx}(-\tau)$



Transmission of a random process through a linear filter

- Suppose RP $X(t)$ is applied to a LTIS with impulse response $h(t)$, producing RP $Y(t)$.
 - What is the mean and autocorrelation of $Y(t)$ w.r.t those of $X(t)$
Assuming $X(t)$ is WSS RP.

$$m_y(t) = E[y(t)] = E[h(t) \otimes X(t)]$$

$$m_y(t) = E\left[\int_{-\infty}^{\infty} h(\tau)X(t-\tau)d\tau\right] = \int_{-\infty}^{\infty} E[h(\tau)X(t-\tau)]d\tau$$

$$m_y(t) = \int_{-\infty}^{\infty} h(\tau)E[X(t-\tau)]d\tau = h(t) \otimes m_x(t)$$

Since $X(t)$ is WSS RP, $m_x(t) = m$

$$m_y(t) = m \int_{-\infty}^{\infty} h(\tau)d\tau = mH(0)$$

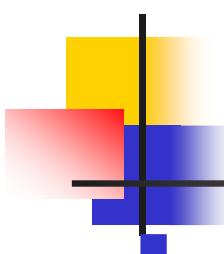
$H(0)$ is zero frequency response of the system

$$\begin{aligned} R_y(t_1, t_2) &= E[y(t_1)y(t_2)] = \\ &= E[(h(t_1) \otimes X(t_1))(h(t_2) \otimes X(t_2))] \\ &= E\left[\int_{-\infty}^{\infty} h(\lambda)X(t_1 - \lambda)d\lambda \int_{-\infty}^{\infty} h(v)X(t_2 - v)dv\right] = \\ &\quad \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(\lambda)h(v)E[X(t_1 - \lambda)X(t_2 - v)]d\lambda dv \end{aligned}$$

$$R_y(t_1, t_2) = h(t) \otimes h(t) \otimes R_x(t_1, t_2)$$

Since $X(t)$ is WSS RP, $R_x(t_1, t_2) = R_x(t_1 - t_2)$

$$R_y(\tau) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(\lambda)h(v)R_x(\tau - \lambda + v)d\lambda dv$$



Power Spectral Density

From Communication Theory we know that Autocorrelation is the IFT of PSD

- IF $X(t)$ is WSS RP

$$R_X(\tau) \xleftrightarrow{F.T} S_X(f)$$

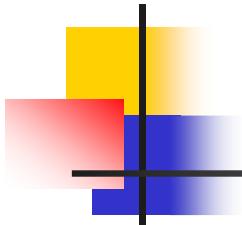
- Properties of PSD

1. $S_X(0) = \int_{-\infty}^{\infty} R_x(\tau) d\tau$

2. $E[X^2(t)] = R_X(0) = \int_{-\infty}^{\infty} S_X(f) df$

3. $S_X(f) \geq 0$ Non Negative

4. $S_X(f) = S_X(-f)$ for real valued random process

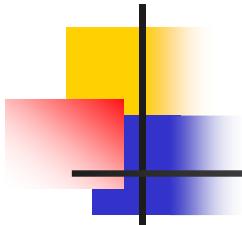


Ex. PSD of sine wave with random phase

$$R_x(\tau) = \frac{A^2}{2} \cos(2\pi f_c(\tau)) \text{ W.S Stationary}$$

- Using the F.T

$$S_x(f) = \frac{A^2}{4} [\delta(f - f_c) + \delta(f + f_c)]$$



Ex.

$$R_x(\tau) = \begin{cases} A^2(1 - \frac{|\tau|}{T}) & |\tau| \leq T \\ 0 & Otherwise \end{cases}$$

- Using the F.T

$$S_x(f) = A^2 T \sin c^2(fT)$$

Ex.

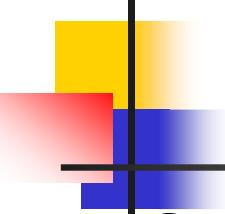
$$Y(t) = X(t) \cos(2\pi f_c t + \theta)$$

$$R_Y(\tau) = E[Y(t)Y(t+\tau)] = E[X(t)\cos(2\pi f_c t + \theta)X(t+\tau)\cos(2\pi f_c(t+\tau) + \theta)]$$

$$R_Y(\tau) = E[X(t)X(t+\tau)]E[\cos(2\pi f_c t + \theta)\cos(2\pi f_c(t+\tau) + \theta)]$$

$$R_Y(\tau) = R_X(\tau) \cos(2\pi f\tau)$$

$$S_Y(f) = \frac{1}{2}[S_X(f - f_c) + S_X(f + f_c)]$$



Relation among PSD of the input and output RP of LTIS

- Suppose RP $X(t)$ is applied to a LTIS with impulse response $h(t)$, producing RP $Y(t)$.
 - What is the PSD of $Y(t)$ w.r.t PSD of $X(t)$ Assuming $X(t)$ is WSS RP.

$$S_Y(f) = S_X(f) |H(f)|^2$$

Ex.

$$H(f) = 1 - \exp(-j2\pi fT)$$

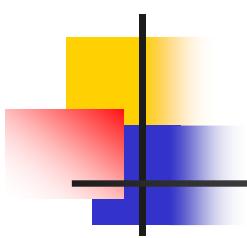
$$|H(f)|^2 = H(f)H^*(f)$$

$$|H(f)|^2 = (1 - \exp(-j2\pi fT))(1 - \exp(+j2\pi fT))$$

$$|H(f)|^2 = (1 + 1 - \exp(-j2\pi fT) - \exp(+j2\pi fT))$$

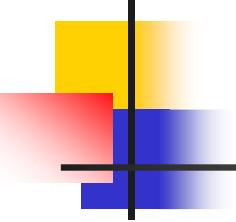
$$|H(f)|^2 = 2(1 - \cos(2\pi fT)) = 4 \sin^2(\pi fT)$$

$$S_Y(f) = 4 \sin^2(\pi fT) S_X(f)$$



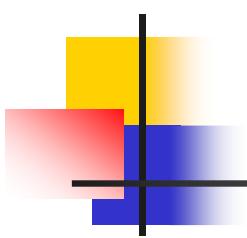
Cross Spectral Density

- Provides a measure of the frequency interrelationship between two random processes.
- $X(t)$ and $Y(t)$ are jointly WSS RP with $R_{XY}(\tau)$ and $R_{YX}(\tau)$
- Thus they have as a FT. $S_{XY}(f)$ and $S_{YX}(f)$
- $S_{XY}(f) = S_{YX}(-f)$



Noise

- Unwanted signal that tend to disturb the transmission and processing of signals in communication systems.
- Thermal Noise → random motion of electrons in a conductor.
- Shot noise → arises in electronic devices, sudden change in voltage or current.



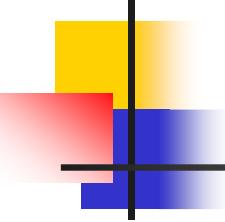
White Noise

- White → occupies all frequencies → PSD is independent on the operating frequency
- Dimensions of N_o is watt per Hertz, $N_o = KT$

$$S_N(f) = \frac{N_o}{2}$$

$$R_N(\tau) = \frac{N_o}{2} \delta(\tau)$$

- Any two different samples of white noise no matter how close they are will be uncorrelated.
- If white noise is Gaussian then they will also be independent



Example

- $X(t)$ and $Y(t)$ have zero mean and they are individually stationary in the wide sense.
- $Z(t) = X(t) + Y(t)$ find $S_Z(f)$

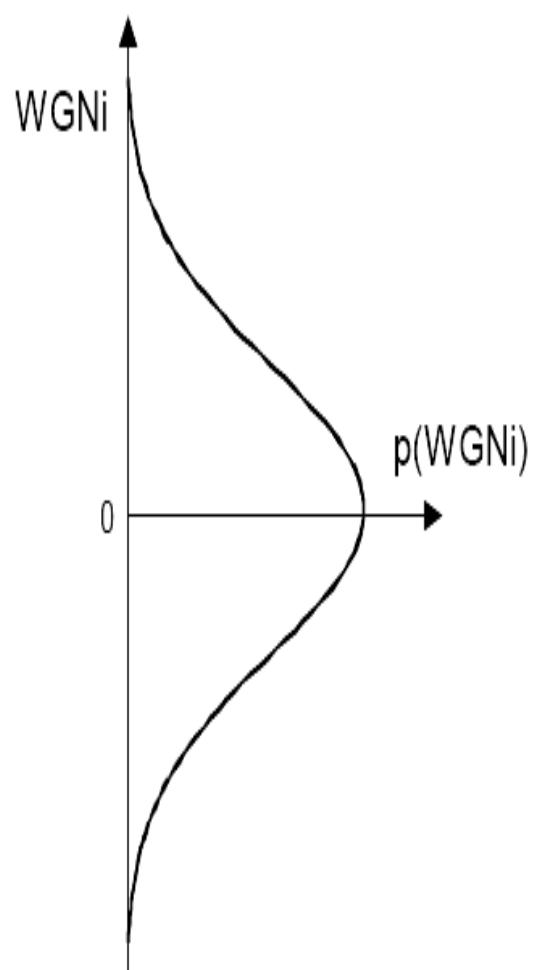
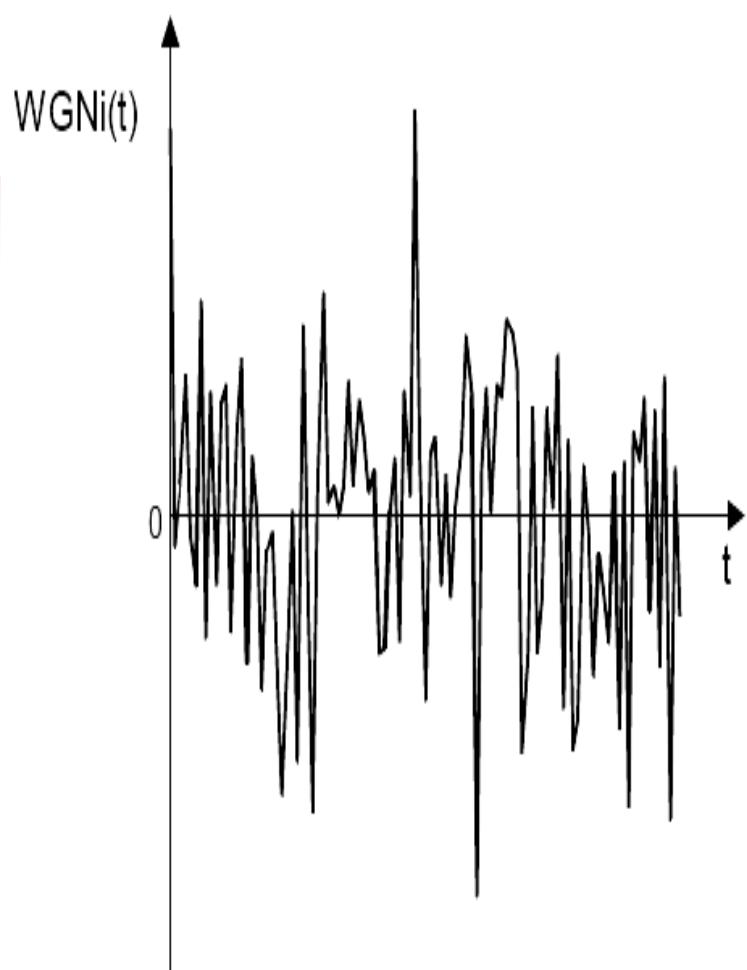
$$\begin{aligned}R_Z(t_1, t_2) &= E[Z(t_1)Z(t_2)] \\&= E[(X(t_1) + Y(t_1))(X(t_2) + Y(t_2))] \\&= E[X(t_1)X(t_2)] + E[X(t_2)Y(t_1)] + E[X(t_1)Y(t_2)] + E[Y(t_1)Y(t_2)]\end{aligned}$$

$$R_Z(\tau) = R_X(\tau) + R_Y(\tau) + R_{XY}(\tau) + R_{YX}(\tau)$$

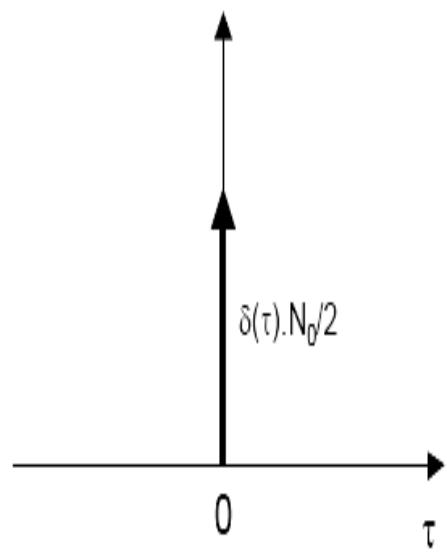
$$S_Z(f) = S_X(f) + S_Y(f) + S_{XY}(f) + S_{YX}(f)$$

If X and Y are uncorrelated

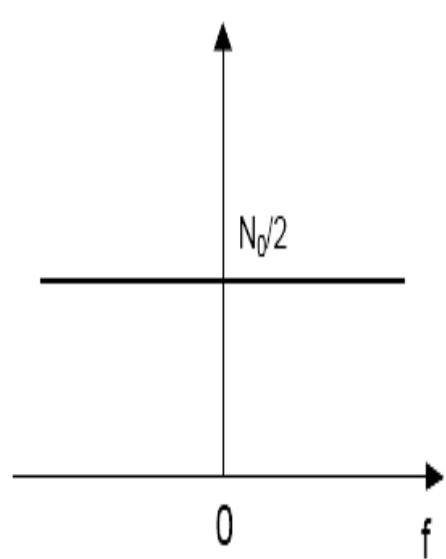
$$S_Z(f) = S_X(f) + S_Y(f)$$

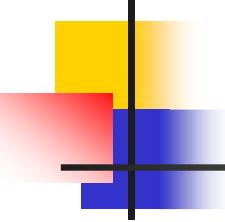


$Ra_{WGN}(\tau)$



$G_{WGN}(f)$



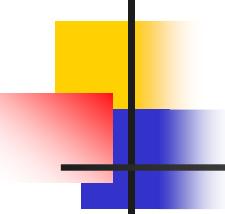


Ideal low pass filtered white noise

- A white Gaussian noise with zero mean and variance $N_0/2$ is applied to an ideal low pass filter of bandwidth B and amplitude response of one.
- PSD of output $Y(t)$ is

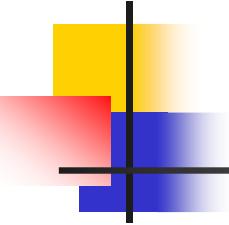
$$S_Y(f) = \begin{cases} \frac{N_o}{2} & -B \leq f \leq B \\ 0 & otherwise \end{cases}$$

- The autocorrelation is $R_Y(\tau) = N_0 B \sin c(2B\tau)$
- Autocorrelation maximum at $\tau = 0$ equal zero equal $N_0 B$ and passes through zero at $\tau = n/2B$ for $n =$ integer values and variance $N_0 B$
- If noise is sampled at rate $2B$ then they are uncorrelated and being Gaussian then statistically independent



RC low pass filtered white noise

- The $H(f)$ of RC filter is
$$H(f) = \frac{1}{1 + j2\pi f RC}$$
- The PSD of the o/p is
$$S_Y(f) = \frac{1}{1 + (2\pi f RC)^2}$$
- The autocorrelation of the output is
$$R_Y(\tau) = \frac{N_o}{4RC} \exp(-\frac{|\tau|}{RC})$$
- If noise is samples at rate $0.217/RC$ then they are uncorrelated and being Gaussian then statistically independent



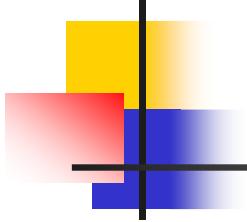
Ex sine wave plus white noise

$$X(t) = A \cos(2\pi f_c t + \theta) + N(t)$$

θ is Uniformely distributed, $N(t)$ is WGN

$$R_X(\tau) = E[X(t)X(t+\tau)] = \frac{A^2}{2} \cos(2\pi f_c \tau) + \frac{N_0}{2} \delta(\tau)$$

$$S_X(f) = \frac{A^2}{4} [\delta(f - f_c) + \delta(f + f_c)] + \frac{N_0}{2}$$



Noise Equivalent Bandwidth

- Output average power of ILPF $\rightarrow N_o B$
- Output average power of RCLPF $\rightarrow N_o / 4RC$

- Output average power of any filter $\rightarrow N_o B H^2(0)$
- Equivalent Bandwidth =
$$B = \frac{\int_0^\infty |H(f)|^2 df}{H^2(0)}$$